

# High-Frequency Trading in the Stock Market and the Costs of Option Market Making

Mahendrarajah Nimalendran <sup>a,\*</sup>, Khaladdin Rzayev <sup>b,c,\*</sup>, Satchit Sagade <sup>d,\*</sup>

<sup>a</sup> *University of Florida*

<sup>b</sup> *University of Edinburgh*

<sup>c</sup> *Koç University*

<sup>d</sup> *Goethe University, Frankfurt*

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## Abstract

Using a comprehensive panel of 2,969,829 stock-day data provided by the Securities and Exchange Commission (MIDAS), we find that HFT activity in the stock market increases market-making costs in the options markets. We consider two potential channels: the *hedging channel* and the *arbitrage channel*. We find that HFTs' liquidity-demanding orders increase the hedging costs due to a higher stock bid-ask spread and a higher price impact for larger hedging demand. The arbitrage channel subjects the option market maker to the risk of trading at stale prices. We show that the hedging (arbitrage) channel is dominant for ATM (ITM) options. Given the significant growth in options trading, we believe that our study highlights the necessity to better understand the costs/risks due to HFT activities in equity markets on derivative markets.

Keywords: market microstructure; high-frequency trading; option market making; hedging; liquidity.

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\*Emails: [mahen.nimalendran@warrington.ufl.edu](mailto:mahen.nimalendran@warrington.ufl.edu), [khaladdin.rzayev@ed.ac.uk](mailto:khaladdin.rzayev@ed.ac.uk), [sagade@safe.uni-frankfurt.de](mailto:sagade@safe.uni-frankfurt.de).

Corresponding author's contact information: University of Edinburgh Business School, University of Edinburgh; e-mail: [khaladdin.rzayev@ed.ac.uk](mailto:khaladdin.rzayev@ed.ac.uk)

We thank Phil Mackintosh and Jeffrey W. Smith at NASDAQ for data provision. The data is made available at no cost to academics following the provision of a project description and the signing of a nondisclosure agreement. Khaladdin gratefully acknowledges the support of the Trans-Atlantic Platform by the Economic and Social Research Council (ESRC) [grant number ES/R004021/1]. This project is funded by the Deutsche Forschungsgemeinschaft (DFG, German Research Foundation) – Projektnummer 329107530.

## 1. Introduction

High-frequency trading (HFT) has materially impacted the dynamics of electronic markets. The large and growing literature examining the implications of HFT has predominantly focused on the within-market, mainly the stock market, quality effects of HFT.<sup>1</sup> Investigations into the cross-asset impact of HFT are limited. In this paper, we attempt to fill this gap by examining the impact of HFT in the stock market on options market microstructure.<sup>2</sup> Figure 1 shows the time-series evolution of trading volume in the US equity and options markets. We see that from 1996 to 2020, the options market volume grew at an annualized compound rate of 15% compared to 11% for the stock trading volume.

### INSERT FIGURE 1 ABOUT HERE

The increase in options trading volume has been observed along with some notable trends and records in the last two years. For instance, for the first time in history, the number of shares traded with options contracts was higher than the underlying stock market trading volume in 2020.<sup>3</sup> Furthermore, trading volume in US equity options markets hits a record high in two consecutive years, 2020<sup>4</sup> and 2021<sup>5</sup>. Given the importance of the options markets in price discovery over time, we believe that studying the potential externalities imposed by HFTs in the stock market on options market is timely and important.

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<sup>1</sup> Hendershott *et al.* (2011), Brogaard *et al.* (2015), Van Kervel and Menkveld (2019) and Hagströmer and Nordén (2013) examine the effects on stock market liquidity; Kirilenko *et al.* (2017) and Lee (2015) investigate the impact in the futures market; Chaboud *et al.* (2014) and Jiang *et al.* (2014) focus on FX and fixed-income securities markets, respectively.

<sup>2</sup> Throughout this paper, we use the phrase high-frequency trading or HFT to refer to HFT activity in the stock market, unless we explicitly indicate otherwise. Furthermore, the acronym HFT is used interchangeably to refer to high-frequency traders and high-frequency trading.

<sup>3</sup> [https://finance.yahoo.com/news/option-trading-volume-higher-underlying-211006236.html?guccounter=1&guce\\_referrer=aHR0cHM6Ly93d3cuZ29vZ2x1LmNvbS8&guce\\_referrer\\_sig=AQAAAEDVmOUhGSccv7vJXLzChzsOdqv\\_dwRYGuoAr4To9lPwq1ho\\_ANZqf8yViK5YWjwDoNZAawTz64F1XrmDdCFkag0FKL5OBmTJ1K0OvgXGlljm\\_wmfjPiDhIEsOjo3HMIO9sghsOBOjYIpvj9KrYsEGRvPPi mhzoNXO1gxEtP0ZKS](https://finance.yahoo.com/news/option-trading-volume-higher-underlying-211006236.html?guccounter=1&guce_referrer=aHR0cHM6Ly93d3cuZ29vZ2x1LmNvbS8&guce_referrer_sig=AQAAAEDVmOUhGSccv7vJXLzChzsOdqv_dwRYGuoAr4To9lPwq1ho_ANZqf8yViK5YWjwDoNZAawTz64F1XrmDdCFkag0FKL5OBmTJ1K0OvgXGlljm_wmfjPiDhIEsOjo3HMIO9sghsOBOjYIpvj9KrYsEGRvPPi mhzoNXO1gxEtP0ZKS)

<sup>4</sup> <https://www.thetradenews.com/occ-clears-record-volumes-for-us-exchange-listed-options-in-2020/>

<sup>5</sup> <https://www.reuters.com/article/us-usa-stocks-options-idUSKBN29K2OI>

In this study, we specifically address the following questions: (i) How does HFT in stocks impact the liquidity of options written on those stocks? (ii) Does the effect vary by option moneyness? (iii) Do different HFT strategies affect option market liquidity differently? (iv) Is the effect exclusively via the stock liquidity channel, or is there a direct effect after controlling for stock liquidity?

We construct HFT measures by using the Securities and Exchange Commission's (SEC) Market Information Data Analytics System (MIDAS) data and show that HFT activity in the equity markets is associated with a significant deterioration in liquidity (increase in quoted bid-ask spreads) in the options markets. Specifically, on average, a one standard deviation increase in HFT activity is associated with a 4.27% (8.89%) higher proportional (dollar) bid-ask spread. To address potential endogeneity concerns between the options market spread and stock market HFT, we employ a two-stage least square (2SLS) instrumental variable (IV) approach by using two different sets of instruments. First, we follow Lee and Watts (2021) and use the randomized experiment of tick size changes launched by the SEC as an instrument for the level of HFT (Hagströmer & Nordén 2013 also employ tick size change as an exogenous shock on HFT). Second, in the spirit of Hasbrouck and Saar (2013), we instrument the level of HFT in a stock-day with the average level of HFT on that day in all other stocks in the same market size quintile. Our results remain robust to both approaches. Moreover, the consistency between the 2SLS IV approach and the fixed effects OLS regression confirms that the relationship between HFT and options market liquidity can be interpreted causally.

We then propose two channels to explain the relationship between HFT and options liquidity: (i) the *hedging* channel and (ii) the *arbitrage* channel. The hedging channel is based on HFT activity in the stock market, affecting the option market bid-ask spread through its effect on option market makers' hedging costs. Black and Scholes (1973) show that option market makers can perfectly hedge their exposures by acquiring an offsetting position in the

underlying asset and continuously rebalancing their portfolio to ensure it remains delta-neutral. However, due to frictions and market imperfections, they can only imperfectly hedge their positions. Consequently, they require compensation for the transaction costs and the risks associated with imperfect hedging of their exposures (see Cho & Engle 1999; Kaul *et al.* 2004; Wu *et al.* 2014).<sup>6</sup>

Boyle and Vorst (1992) and Cho and Engle (1999) demonstrate that option market makers' hedging costs are proportional to the stock market bid-ask spread (see also Kaul *et al.* 2004). This implies that the hedging channel leads to two opposing predictions for the impact of HFT on the option spread. On the one hand, option market makers can more precisely hedge their exposures at lower costs if liquidity in the underlying stock market improves. This is likely when liquidity supplying HFT activity is high, as argued by Hendershott *et al.* (2011) and Brogaard *et al.* (2015). On the other hand, if liquidity is lower in the underlying stock market, option market makers will likely quote wider spreads due to increased costs of hedging their exposure and keeping their positions partially unhedged by reducing the rebalancing frequency. Arbitrage and momentum/directional strategies employed by HFT firms rely on aggressive trades<sup>6</sup> and are harmful for overall market liquidity (see Budish *et al.* 2015; Foucault *et al.* 2016; Foucault *et al.* 2017).

Kaul *et al.* (2004) and Engle and Neri (2010) show the transaction costs of hedging are due to setting up and unwinding the initial hedge and rebalancing costs. At-the-money (ATM) options have high rebalancing costs due to their high gamma. In-the-money (ITM) options have the highest absolute delta and high costs of initially setting up and unwinding the hedge. By contrast, out-of-the-money (OTM) options have the lowest delta and gamma, and therefore, the hedging cost is relatively lower in these options. Above discussion implies that, if the hedging

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<sup>6</sup> Leland (1985) and Boyle and Vorst (1992) develop alternative discrete-time option replication strategies in the presence of transaction costs.

channel indeed explains the association between HFT and the cost of option market making, then the impact of HFT activity in the underlying market on option spreads should be higher for ATM and ITM options.

The arbitrage channel relates to violations of the put-call parity relationship from the asynchronous adjustment in stock and option prices. These “toxic” arbitrage opportunities, to the extent driven by a delay in incorporating information into stock and the option prices, will induce market-makers to post wider quotes to protect against adverse selection losses (see Budish *et al.* 2015; Foucault *et al.* 2017). The short-lived nature of such arbitrage opportunities requires arbitrageurs to rely on aggressive orders in the stock (see Kozhan & Tham 2012). Options market-makers are especially exposed to such toxic arbitrage losses due to the exchange-imposed caps on the number of quote updates and fines on traders with higher messages-to-transactions ratios (see Muravyev & Pearson 2020). These restrictions severely limit the option market makers’ ability to update their quotes in response to new information. In addition, liquidity-consuming HFTs engaging in cross-market arbitrage strategies may exploit violations of the put-call parity relation by sniping stale quotes in the options market. Hence, aggressive HFTs can expose option market makers to the options market risk of trading at stale quotes. Notably, Halpern and Turnbull (1985) and Galai (1978) observe that violations of the put-call parity relationship are more frequent for ITM options. Thus, if the arbitrage channel is the dominant channel to explain the relationship between HFT and the costs of option market making, then the impact of HFT on the option spread should be higher for ITM options and should be weakened on days without information.

To test the hedging and arbitrage channels, we first examine whether the effects of HFT on option spreads vary by three moneyness groups: (i) ATM, (ii) ITM, and (iii) OTM options. While the results are statistically significant in all moneyness groups, consistent with the arbitrage and hedging channels, the economic magnitudes of the impact are relatively higher

for ATM and ITM options. However, it is important to note that our data (the SEC's MIDAS data) does not allow us to directly test the hedging and arbitrage channels. Specifically, the direct testing of the hedging and arbitrage channels requires us to use more granular HFT data; one must disaggregate HFT activity into aggressive and passive trades to test these channels. The MIDAS data allows us to construct only a general measure of HFT from this data. Therefore, we employ proprietary HFT data provided by NASDAQ to further test the hedging and arbitrage channels. Unfortunately, this data is only available for 2009; however, due to its granularity, it complements the MIDAS data that covers a much longer period, 2012-2019.

We conduct two additional tests using the NASDAQ's proprietary data. In the first analysis, we find that after controlling for the stock market bid-ask spread, the impact of liquidity consuming HFT activity on option spread remains positive and significant for ATM and ITM options. Specifically, a one standard deviation increase in aggressive HFT activity increases ATM option proportional (dollar) spreads by about 5.71% (13.89%).<sup>7</sup> For ITM options, a one standard deviation increase in aggressive HFT activity is associated with a +6.22% (+9.66%) change in the proportion (dollar) option market spread. Importantly, the association between aggressive HFT activity and ITM option spreads is weakly statistically significant, while the economic magnitude is quite high. Finally, the results are insignificant for OTM options. These findings are consistent with the predictions of the hedging and arbitrage channels.

In the second analysis, we further isolate the arbitrage channel by identifying and excluding days involving the release of firm-specific news from our sample. We observe that the impact of aggressive HFT on ITM option spread becomes insignificant. On the other hand, for ATM options, while the magnitude of the effect decreases by about 19% for proportional

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<sup>7</sup> The estimates based on the NASDAQ proprietary data is similar to the estimates we obtain from the longer panel based on the SEC-MIDAS data (4.27% (8.89%) for the proportional and (dollar) spreads).

spread, the relationship remains significant. These two tests indicate that the hedging (arbitrage) channel predominantly explains the relationship between aggressive HFTs, and options market spread for ATM (ITM) options. However, we caution against over-interpretation of this particular result because the NASDAQ HFT data has two important limitations. First, it covers only 103 randomly selected stocks for the year 2009. Second, the data includes HFT activity in the NASDAQ only. Furthermore, we find that the economic magnitude of the impact of HFT on option spreads is substantial in ITM contracts when we use the MIDAS HFT data, which covers most of the US-listed common stocks and the period of 2012 to 2019. Therefore, we believe that the association between HFT and ITM option spreads can be likely due to both hedging and arbitrage channels.

Interestingly, the effect of liquidity providing HFT activity on option spread is not statistically significant. We show that this is because 100% of passive HFT trades' effect on options market spread is captured by the bid-ask spread in the stock market. Conversely, 75% of the impact of aggressive HFT activity on spreads in the options market is direct and unrelated to the stock market bid-ask spread. This is likely due to several factors: the bid-ask spread capturing transaction costs for small(er) orders due to the limited depth available at the best quotes, the predictability of option market makers' hedging demand, option market makers employing order splitting algorithms to execute large hedging trades, and aggressive HFTs engaging in predatory trading (see Brunnermeier & Pedersen 2005).

Our findings offer insights into the role of HFTs in explaining the cross-asset market microstructure dynamics between the stock market and the options market. To our knowledge, our study is the first to provide evidence on the direct impact of various strategies used by HFTs on the option spread. In a recent study, Kapadia and Linn (2019) use the Glosten and Milgrom (1985) framework to develop a model of trading in the primary market (stock) and a derivative market (option). The model is used to relate the bid-ask spread in the options market

to the volatility of the bid-ask spread in the stock market via a synthetic stock based on put-call parity. The authors use a glitch in Knight Capital's trading platform that erroneously executed a large number of small orders for a selected number of stocks. While the glitch increased uninformed order flow, it also resulted in persistent liquidity-related uncertainty. The study finds option spreads widened among impacted stocks and remained wide for a quarter of an hour after the broker-dealer had fixed the glitch in their computer system.

Our study also examines how low-latency traders (HFTs) in the stock market on the option spreads. While Kapadia and Linn (2019) analyze the effects of stock market quote uncertainty on the options market spread, we focus directly on the impact of HFTs on option spreads by using comprehensive HFT data. Our study, owing to the rich and granular MIDAS and NASDAQ HFT datasets, provides direct evidence on the negative impact of HFT on the options market making. While HFTs may cause uncertainty in the underlying stock liquidity, it is not solely driven by HFTs. Thus, the uncertainty in liquidity is not necessarily a proxy for HFT activity and is not commonly used in the literature. Nevertheless, for robustness, we control for the volatility of the stock bid-ask spread by including it as an explanatory variable and obtain the same results (see Table A.1).

The focus of our study is also related to the work of Mishra *et al.* (2012), one of the first studies to examine the impact of automation on options markets. Mishra *et al.* (2012) use high-frequency data sourced from the Options Price Reporting Authority (OPRA), and show that automation reduces bid-ask spreads and increases liquidity in options markets. Our study differs from theirs in at least two ways. First, they focus on the impact of automated trading on option liquidity. While HFT strategies require automated trading, HFTs are only a subset of all the traders on the automated trading platform, and they have different impacts on market dynamics. Secondly, and most importantly, Mishra *et al.* (2012) investigate the relation between option liquidity and option automated trading. However, as discussed, liquidity in



options markets is also determined by equity market dynamics (see Cho & Engle 1999). Since HFTs are one of the most important drivers of the equity market liquidity, it is expected that the option spread will also be impacted by HFT activity in the underlying market.

Although HFTs use both market-making and speculative trading strategies in the stock market, it is widely accepted that the majority of them (about 80%) follow the market-making strategy, implying that the adverse-selection-avoidance channel dominates the picking-off channel in this market (see Hagströmer & Nordén 2013; Menkveld 2013). In this context, if HFTs predominantly engage in market making in the options market, too, then our results suggest that liquidity-demanding HFTs in the stock market will increase the costs of liquidity-supplying HFTs in the options market. As such, this study provides a complementary perspective to that of Menkveld and Zoican (2017), who show that speculative HFTs impose adverse-selection cost on market-making HFTs. Moreover, focusing on the stock and option markets, we find that this is also the case in the cross-market setting.

In other related literature, studies examine the effects of various market microstructure determinants of the option spread. Easley *et al.* (1998) propose a model where informed traders choose between stock and option markets based on the relative transaction costs in the markets and the “bang-for-buck” in the form of leverage afforded by the options market. The authors conclude that depending on the relative transaction costs in the markets there can be a separating equilibrium where informed traders trade only in the stock market or a pooling equilibrium where informed traders trade in both markets. Subsequent empirical work that focused on informed trading in these two venues has largely supported the model’s theoretical predictions. For example, Cao *et al.* (2005) find informed options trading before takeovers. Hu

(2014) provides evidence of an information channel by documenting that the option market makers' initial delta hedging strategy is reflected in stock prices.<sup>8</sup>

*Roadmap:* The rest of this paper is organized as follows. In Section 2, a description of the data is provided. Section 3 presents the main results, and Section 4 gives the conclusion.

## 2. Data and Variable Construction

### 2.1. Data Sources

We compile data from several sources. The Securities and Exchange Commission's Market Information Data Analytics System (MIDAS) data is used to construct HFT proxies. The SEC offers MIDAS data to promote investigations on the US equity markets structure. The MIDAS collects data across all major US exchanges since 2012, and it includes the following variables: lit trade and order volume, hidden trade and order volume, odd lot trade and order volume, and counts of cancellations (full or partial) for each day. The dataset covers over 5,570 stocks and 2,730 exchange-traded funds.

End-of-the-day option bid and ask prices, trading volumes, Greeks, and implied volatilities are obtained from OptionMetrics. Greeks and implied volatilities are computed by using a binomial tree where an interest rate is constant. We obtain options data for all US-listed common stocks from 2012 to 2019 as MIDAS data is available from 2012 only. Our sample does not include 2020 due to the Covid-19 induced volatility. We follow the existing literature and exclude long-term options, i.e., those with maturities greater than 180 days. This allows us to restrict our analysis to the most actively traded options contracts, i.e., options contracts with higher trading volume (see Brenner *et al.* 2001; Christoffersen *et al.* 2017). There are over 1

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<sup>8</sup> Pan and Poteshman (2006), Ni *et al.* (2008), Cremers and Weinbaum (2010), Ge *et al.* (2016) and Collin-Dufresne *et al.* (2020) document the role of options markets in the price discovery process, that is consistent with the information channel.

billion option transactions in the final sample, and the total nominal and USD volumes of these transactions are approximately 28.7 billion and US\$73.9 billion, respectively.

Daily ask, bid, and trading prices are obtained from the Center for Research in Securities Prices (CRSP) dataset. The main analysis includes all CRSP common stocks matched in the MIDAS and OptionMetrics databases. The resulting sample has 2,746 unique securities and 2,969,829 security-day observations.

## 2.2. Variable Construction

We construct five HFT measures from the MIDAS data. Our first proxy is the ratio of the quote to trade volume (see Hendershott *et al.* 2011),  $QT_{i,d}$ , which is computed as the sum of order volume for all order messages divided by the sum of trade volume for all trades that are not against hidden orders. The second HFT measure is the cancel-to-trade ratio,  $CT_{i,d}$ , or the number of all cancel messages (full or partial) divided the number of trades (see Weller 2018). The third (fourth) HFT proxy is the odd lot rate (odd-lot volume),  $OR_{i,d}$  ( $OV_{i,d}$ ), which is calculated as the number of odd lot trade messages (odd lot trade volume) divided by the number of all trade messages (trade volume) (see O'Hara *et al.* 2014).<sup>9</sup> The final HFT measure is the inverse of the average trade size,  $ITS_{i,d}$ , or the number of trades divided by the number of shares traded (see Conrad *et al.* 2015).

Consistent with the literature, we measure option liquidity using bid-ask spreads (see as an example, Muravyev & Pearson 2020). Our liquidity proxies are the proportional quoted spread, computed as the difference between the ask and bid prices divided by the midpoint and the dollar quoted spread, which equals the difference between the ask and bid prices for each options transaction. Options contracts with various characteristics (different maturities and strike prices) are traded for each stock and day. As a result, for each stock, we observe multiple

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<sup>9</sup> Odd lots are trades that has a volume of less than 100 shares.

bid-ask spreads during the day. Given that our main analysis is based on panel regressions at a daily frequency for each underlying, we follow Muravyev and Pearson (2020) and compute the daily proportional spread ( $OPspread_{i,d}$ ) and dollar spread ( $ODspread_{i,d}$ ) as the dollar-volume-weighted average of all spreads for stock  $i$  and day  $d$ .

Apart from the variables mentioned above, we employ several control variables to capture stock and option market dynamics. Our option market variables are the options volume ( $Ovolume_{i,d}$ ), implied volatility ( $Oimplied_{i,d}$ ), absolute option delta ( $|Odelta_{i,d}|$ ), option vega ( $Ovega_{i,d}$ ), and option gamma ( $Ogamma_{i,d}$ ). The  $Ovolume_{i,d}$  is the natural logarithm of the daily trading volume (contracts) for each stock  $i$  and day  $d$ . OptionMetrics provide the implied volatility and option Greeks. Similar to option market spread measures, we compute the daily  $Oimplied_{i,d}$ ,  $|Odelta_{i,d}|$ ,  $Ovega_{i,d}$ , and  $Ogamma_{i,d}$  as the dollar-volume-weighted averages of all implied volatilities, absolute deltas, vegas, and gammas for stock  $i$  and day  $d$ . We use the absolute value of delta as call and put options have different signs for deltas, i.e., a call option delta is positive while delta is negative for put options.

We employ the following variables to control for stock market activity: stock spreads ( $SPspread_{i,d}$  and  $SDspread_{i,d}$ ), stock price volatility ( $SVolatility_{i,d}$ ).  $SPspread_{i,d}$  (proportional) and  $SDspread_{i,d}$  (dollar), are computed using the best ask and bid prices for each stock and day. Specifically,  $SPspread_{i,d}$  is computed as the difference between the best ask and bid prices for stock  $i$  and day  $d$ , divided by the midpoint of the two prices on the same day.  $SDspread_{i,d}$  is the difference between the best ask and bid prices for stock  $i$  and day  $d$ .  $SVolatility_{i,d}$  is the absolute difference between the last transaction prices for stock  $i$  on days  $d$  and  $d-1$ .

### **INSERT TABLE 1 ABOUT HERE**

Table 1 provides an overview of all the variables and their computation methods.

### 2.3. Descriptive Statistics

Table 2 provides descriptive statistics for the 2,746 stocks in the full sample and 1,235 stocks in the tick size pilot sample and their listed options (see Section 3.2 for detailed information about the tick size pilot sample). We winsorized all variables at the 1<sup>st</sup> and 99<sup>th</sup> percentile values.

#### INSERT TABLE 2 ABOUT HERE

Panel A reports the summary statistics for the underlying stock market variables. For the full sample, the average  $SPspread_{i,d}$  ( $SDspread_{i,d}$ ) is 0.11% (0.02 USD), which is lower than the average  $OPspread_{i,d}$  ( $ODspread_{i,d}$ ) by a factor of approximately 310 (22.5). Furthermore,  $SPspread_{i,d}$  and  $SDspread_{i,d}$  are higher for the tick size pilot sample. This is expected as relatively small and illiquid stocks have been included in the Tick Size Pilot Programme (see Chung *et al.* 2020).

Panel B reports the summary statistics for the options market variables for the full sample and the three moneyness groups (ATM, OTM, and ITM) based on the classification provided by Bollen and Whaley (2004). Specifically, we define OTM options as those with absolute option delta  $|\Delta| \leq 0.375$ , ATM options as those with  $0.375 < |\Delta| \leq 0.625$ , and ITM options as those with  $|\Delta| > 0.625$ . For ATM options, the average  $OPspread_{i,d}$  and  $ODspread_{i,d}$  are 21.64% and 0.38 USD, respectively.  $ODspread_{i,d}$  increases as we move from OTM (0.27 USD) to ITM (0.62 USD) contracts. This is not surprising as ITM (OTM) contracts have higher (lower) prices. Wei and Zheng (2010) also show that the dollar spread is higher for options with higher prices. By contrast,  $OPspread_{i,d}$  increases as we move from ITM (15.73%) to OTM (52.10%) contracts. Wei and Zheng (2010) link this to leverage; specifically, they argue that option contracts with higher leverage (OTM contracts) attract more informed traders and are associated with a higher spread. ATM options have the highest  $Ovolume_{i,d}$ ,  $Ovega_{i,d}$  and  $Ogamme_{i,d}$ . This is expected as these contracts are the most active.

Consistent with the equity market spreads, the options spreads are higher for the tick size pilot sample.

#### 2.4. Correlation Between Control Variables and Dependent Variables

We use several variables as controls for the stock and options spread regression models. The option market variables are  $Ovolume_{i,d}$ ,  $Oimplied_{i,d}$ ,  $|Odelta_{i,d}|$ ,  $Ovega_{i,d}$ ,  $Ogamma_{i,d}$  and the stock market variables are  $Sspread_{i,d}$  and  $SVolatility_{i,d}$ . To show that the selected control variables are indeed valid and capture the important variation in option spreads, we first estimate the association between these control variables and option spreads.

#### INSERT TABLE 3 ABOUT HERE

Table 3 reports the correlation between the control variables and dependent variables used in the study. The estimates suggest that the associations between control variables and option spreads are significant and in line with the relevant literature. For instance,  $Ovolume_{i,d}$  is negatively correlated with option spreads suggesting that a higher trading volume implies higher liquidity. Consistent with Cho and Engle (1999) and Engle and Neri (2010), we also find that the stock spread is positively and significantly correlated to the option spread, implying that the stock spread is a significant determinant of the options markets liquidity.

### 3. Estimation Approaches

#### 3.1. Fixed Effect Estimation

We begin testing the impact of HFT on the option spreads by estimating the following fixed-effect models:

$$OSpread_{i,d} = \alpha_i + \beta_d + \gamma_1 HFT_{i,d} + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d} \quad (1)$$

where  $OSpread_{i,d}$  corresponds to either the proportional spread ( $OPspread_{i,d}$ ) or the dollar spread ( $ODspread_{i,d}$ ),  $HFT_{i,d}$  corresponds to one of the five HFT proxies ( $QT_{i,d}$ ,  $CT_{i,d}$ ,  $OR_{i,d}$ ,

$OV_{i,d}$ ,  $ITS_{i,d}$ ). The  $\alpha_i$  and  $\beta_d$  are stock and time (day) fixed effects. Standard errors are double clustered on stock and day.<sup>10</sup> The  $C_{k,i,d}$  is a set of  $k$  control variables, including variables from both the option and underlying markets. The option market variables are  $Ovolume_{i,d}$ ,  $Oimplied_{i,d}$ ,  $|Odelta_{i,d}|$ ,  $Ogamme_{i,d}$ , and  $Ovega_{i,d}$  and the stock market variables are  $SPspread_{i,d}$  (when we employ  $OPspread_{i,d}$  as the dependent variable),  $SDspread_{i,d}$  (when we use  $ODspread_{i,d}$  as the dependent variable), and  $SVolatility_{i,d}$ . All these variables are defined in Table 1.

### 3.2. Two Stages Least Square (2SLS) Instrumental Variable (IV) Approach

The presence of HFT in the stock markets and option bid-ask spreads could be determined by unobserved common factors, implying they are jointly endogenous. To address these potential endogeneity concerns, we estimate the following 2SLS IV approach:

$$OSpread_{i,d} = \alpha_i + \beta_d + \gamma_1 \widehat{HFT}_{i,d} + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d} \quad (2)$$

$$HFT_{i,d} = \alpha_i + \beta_d + \vartheta_1 IV_{i,d} + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d} \quad (3)$$

where  $\widehat{HFT}_{i,d}$  is the fitted values of  $HFT_{i,d}$  obtained by regressing  $HFT_{i,d}$  on  $IV_{i,d}$ . We use two different sets of instruments for robustness. Our first instrument is based on tick size changes (see Hagströmer & Nordén 2013). Specifically, we follow Lee and Watts (2021) and employ the introduction of the Tick Size Pilot Programme as an exogenous shock on HFT. In October 2016, the SEC launched a two-year pilot program to investigate the effect of increased tick size on market quality. The pilot program consists of treatment and control groups of 1,200 randomly selected stocks each. The treatment securities are split into three treatment groups; each group is subject to different changes. The stocks that are included in the first treatment group must be quoted in 5 cent increments. In addition to being subject to the increments in

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<sup>10</sup> Double-clustered standard errors are used in all subsequent models.

tick size applying to the first group, the stocks in the second treatment group must be traded in 5 cent increments. The stocks in the third treatment group have the same quoting and trading rules as the second treatment group. However, they are also subject to the “trade-at” rule. More specifically, orders must not be executed in dark venues for these stocks unless there is a meaningfully better price in dark markets. The tick size sample includes all stocks in the Tick Size Pilot Experiment matched in the MIDAS and OptionMetrics databases. The resulting sample has 1,235 unique securities (617 control and 618 treated stocks) and 640,306 security-day observations.

We follow Lee and Watts (2021) and Chung *et al.* (2020) and use these three groups as our treatment group. The Tick Size Pilot Programme commenced gradually on October 3, 2016, and all treatment firms are included by the end of October; the program was implemented over the next two years. Thus, for the treatment stocks,  $IV_{i,d}$  is a dummy variable that takes the value of one after the change (from October 3, 2016, to September 28, 2018) and zero before (from October 1, 2014, to October 2, 2016). For the control stocks,  $IV_{i,d}$  equals zero during the sample period (from October 1, 2014, to September 28, 2018). Given that we include matched sample of stocks into the first-stage model, the estimation results of the second-stage model effectively give us a difference-in-difference estimate (see Malceniace *et al.* 2019).

While the changes in tick size is used as an exogenous shock on HFT in the relevant market microstructure literature (see Hagströmer & Nordén 2013; Lee & Watts 2021), it is not a perfect setting. The main concern with this identification strategy is that the tick size changes do not influence only the amount of HFT. For instance, Albuquerque *et al.* (2020) show that an increase in tick size led to increased transaction costs (see also Chung *et al.* 2020). Thus, the changes in tick size in equity markets can affect options spreads through its direct impact on the transaction costs in the equity markets. Although we control the stock spread in our first and second stage models to address this issue, we employ an additional instrument as our



second identification strategy for robustness. Specifically, we follow Hasbrouck and Saar (2013) and instrument, the level of HFT in a stock-day with the average level of HFT on that day in all other stocks in the corresponding size quintile (see also Comerton-Forde & Putniņš 2015). This variable meets the requirements for an instrument because of two reasons. First, the level of HFT in other stocks is correlated with the level of HFT in a particular stock. Second, HFT in other stocks is unlikely to be driven by the nature of liquidity in options on a particular stock. This instrument alone is sufficient for identification; however, we follow Foley and Putniņš (2016) and include the first lag values of the dependent variable (HFT measures) as an additional control variable in the first-stage regression.

## **4. Estimation Results**

### **4.1. The Impact of HFT on Option Spreads**

The estimation results of Equations 1 and 2 are presented in Table 4.

#### **INSERT TABLE 4 ABOUT HERE**

The first-stage results of the 2SLS IV approaches are reported in the Appendix section (see Table A.2). Overall, our selected instruments are significantly correlated to all HFT measures, and the signs of the associations are as expected. We scale all independent variables by their standard deviation as this allows us to directly estimate the economic significances of the effect (see Foucault & Fresard 2014). Column 1 presents the results of the standard OLS approach with stock and day fixed effects, while Columns 2 and 3 report the results for the 2SLS IV approach. Overall, we see that HFT in the underlying markets is positively related to the option spreads. This suggests a negative role of HFTs in the options market liquidity. The results are remarkably consistent across various HFT proxies and specifications. Consistency between the OLS and 2SLS IV methods allows us to establish a causal link between HFT and the option spread while ruling out endogeneity concerns.

The magnitude of the impact is economically significant too. For example, estimates using the standard OLS approach with stock and day fixed effects show that, on average, a one-standard-deviation increase in HFT increases  $ODspread_{i,t}$  and  $OPspread_{i,t}$  by 8.89% and 4.27%. This means that for the average option,  $ODspread_{i,t}$  ( $OPspread_{i,t}$ ) will increase from 0.45 USD (34.97%) to 0.49 USD (36.46%). The economic magnitudes of the 2SLS-IV estimates depend on which set of instruments is used. When we use the average HFT of the other stocks in the same size quintile as our instrument in the 2SLS-IV, the economic effects are even larger in magnitude. Specifically, a one-standard-deviation shock on HFT raises  $ODspread_{i,t}$  and  $OPspread_{i,t}$  by 14.22% and 6.46%. However, when we use the changes in tick size as an instrument, a one-standard-deviation increase in HFT is associated with +3.30% and +2.98% change in  $ODspread_{i,t}$  and  $OPspread_{i,t}$ , respectively. As seen, tick size results are lower than those of other specifications. This is because only small stocks have been included in the Tick Size Pilot Programme. This is a plausible explanation as Brogaard *et al.* (2014) show that HFTs are less active in small stocks.

In this paper, we employ five HFT measures and three different model specifications to investigate the relation between HFT and options spread. Hence, we believe that these estimations are sufficient to isolate the effects of HFT on option spreads. Nevertheless, the results can still be driven by some common factors that drive both HFT and option spreads. To rule out this concern, we include several control variables in Equation (2).

In addition to this, we want to specifically discuss one of these common factors, namely stock market volatility. The stock market volatility is one of the determinants of HFT trading volume (see Brogaard *et al.* 2014). Furthermore, as reported in Table 3, volatility is an important determinant of options liquidity. The implication is that our model may capture the impact of stock market volatility on both HFT and options spreads rather than the association between HFT and options spread. While we are controlling for absolute price changes, one of

the common volatility measures in the literature (see Karpoff 1987), in Equation (2), we conduct some moderation tests to further strengthen the interpretations of our results. Specifically, we estimate two additional models. We first test the association between HFT and stock realized volatility (where the realized spread is measured as the variance estimates based on the returns calculated using the midpoint of the quoted bid and ask prices at every second during the trading hours). Second, we examine the association between HFT and options spread after controlling for both the absolute value of price changes and the realized volatility. The estimation results are reported in the Appendix section (Tables A.3 and A.4). Two points stand out. First, although the association between HFT and the realized spread is statistically significant, the signs of the association are not consistent across different HFT measures and specifications. Secondly and importantly, the effects of HFT on options spread are statistically significant even after controlling for the realized volatility. These results suggest that the volatility does not drive our results.

The results provided in Table 4 shows the average effects (across firms and years) of HFT on options spread, suggesting that the estimations can mask time variation in the effects. Hence, to further understand the time variation in the association between HFT and options spread, we estimate Equation (2) year-by-year.

#### **INSERT FIGURE 2 ABOUT HERE**

Figure 2 displays the yearly estimates of  $\gamma_1$ . For  $ODspread_{i,t}$  ( $OPspread_{i,t}$ ), the associations between HFT (measured by  $QT_{i,t}$ ) and options spread are significant in seven (six) of eight years. The results suggest that, overall, the effects of HFT on options spreads is persistent across years.

#### 4.2. The Impact of HFT on Option Spreads by Moneyness

To further study the impact of HFT on option spreads and understand the channels through which HFT impacts option spreads, we split our sample into moneyness groups and estimate Equations 2 and 3 for each group.

##### INSERT TABLE 5 ABOUT HERE

Table 5 presents the estimation results for ATM, ITM, and OTM options, respectively. Similar to the results reported in Table 4, there is a positive and significant relation between HFT and option spreads across the three moneyness groups in most specifications. There are few exceptions for the OTM options. Specifically, for two HFT proxies ( $OV_{i,t}$  and  $TS_{i,t}$ ), the impact of HFT on  $OPspread_{i,t}$  is negative when we use the standard OLS approach. Nevertheless, the results show that the impact of HFT on the options market making is robust and persistent. However, in addition to the statistical significance of the results, we also study the economic significance among different moneyness groups.

##### INSERT TABLE 6 ABOUT HERE

Table 6 presents the economic magnitudes of the effect of HFT on option spreads for the full sample and three moneyness groups. Two observations stand out. First, as discussed above, the magnitude of the increase in option spreads is economically meaningful. Secondly and more importantly, the magnitudes of the impact of HFT on option spreads differ by the moneyness of the options contract. More explicitly, in all specifications, the economic magnitude is higher for ITM (17.10% for  $ODspread_{i,t}$  and 6.80% for  $OPspread_{i,t}$ ) and ATM (6.32% for  $ODspread_{i,t}$  and 7.55% for  $OPspread_{i,t}$ ) options in comparison to OTM options (3.93% for  $ODspread_{i,t}$  and 0.91% for  $OPspread_{i,t}$ ). We conjecture that this finding is important as it may give us insights into the channels that drive the HFT-option market-making relationship.

As suppliers of immediacy, market makers in equity markets face adverse selection, inventory holding, and order-processing costs. Cho and Engle (1999) and Kaul *et al.* (2004) argue that, in addition to adverse selection, inventory holding, and order-processing costs, hedging costs play an important role in determining option spreads. Battalio and Schultz (2011) show that option market makers' exposure to adverse selection and inventory risks is typically larger. First, the option market makers' inventory positions can be highly volatile due to the implicit leverage in options contracts and uncertainty about stock return volatility (see also Jameson & Wilhelm 1992). Second, the option market makers have limited control over their inventory positions due to option market dynamics. For example, traders are more likely to write call options than buy them, while they use buy and sell orders roughly evenly in equity markets (see Lakonishok *et al.* 2007). Due to these reasons, options market makers hedge their inventories by taking an offsetting position in the underlying cash market.

In a discrete-time setting, option market makers' hedging costs consist of two components: the cost of setting up and liquidating the initial delta-neutral position and the cost of continuously rebalancing the portfolio and maintaining a delta-neutral position (see Jameson & Wilhelm 1992; Cho & Engle 1999; Kaul *et al.* 2004; Engle & Neri 2010). Kaul *et al.* (2004) show that to the extent option market makers employ market orders to hedge their inventories, their hedging costs are proportional to the stock market bid-ask spread. In addition to the bid-ask spread, the rebalancing component of the hedging costs is positively related to the volatility of the underlying asset and the sensitivity of the option to changes in underlying volatility (*option vega*,  $v$ ) inversely related to the revision interval.

The magnitudes of the option market makers' hedging costs differ by the moneyness of the options contract. For example, ITM options have the highest absolute delta and hence have the highest cost associated with the setup and unwinding of the initial hedge. On the other hand, ATM options have the highest gamma and vega, and hence the market-makers positions in

these contracts need to be rebalanced/hedged much more frequently (see Kaul *et al.* 2004). Consequently, ATM options have higher rebalancing costs than ITM and OTM options, and ITM options have higher setup and unwinding costs of the delta-neutral position than ATM and OTM options (see Wu *et al.* 2014). This implies that OTM options have the lowest initial and rebalancing hedging costs. Noticeably, as reported in Table 6, we also find that the economic magnitudes of the impact of HFT on option spreads are the lowest for the OTM contracts, suggesting that the *hedging* channel is a plausible explanation for our findings of the impact of HFTs on option spreads.

The hedging channel discussed above may not be the only channel through which HFT activity in the stock market affects option market spreads. The payoffs of a stock and its listed option contracts are correlated through put-call parity. This parity relationship states that a portfolio that consists of a short put option and a long call option – both with the same strike price and maturity date – will have the same return as holding a forward contract with the same strike price and maturity. If this relationship does not hold, there will be a violation of the law of one price, implying the existence of an arbitrage opportunity between the stock and option markets market (see Galai 1978; Halpern & Turnbull 1985; Ofek *et al.* 2004).<sup>11</sup>

Options market-makers are particularly exposed to such toxic arbitrage losses as the exchange-imposed caps on the number of quote updates and fines on traders with higher messages-to-transactions ratios (see Muravyev & Pearson 2020) severely limit the option market makers' ability to update their quotes in response to new information. Therefore, liquidity-consuming HFTs engaging in cross-market arbitrage strategies may exploit the put-call parity relation violations by sniping stale quotes in the options market. Halpern and Turnbull (1985) and Galai (1978) observe that the frequency of profitable put-call parity

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<sup>11</sup> In case of American options, the above-mentioned put-call parity relation is typically expressed as an inequality due to the early-exercise premium of American calls and puts.

violations in ITM contracts is high due to the prices of such options following stock prices very closely and the resulting difficulty in keeping the option and stock prices arbitrage-free in these contracts. Consistent with these findings, our results also suggest that the impact of HFT on option spread (especially for the dollar spread) is relatively higher for ITM options. Thus, the *arbitrage* channel might also be a potential channel to explain the association between HFT and option spreads.

The main implication of the hedging and arbitrage channels is that they predict heterogeneity in the impact of various HFT strategies (liquidity-supplying and -demanding) on option spreads. In the hedging channel, HFT activity in the stock market likely affects option market makers' hedging costs by affecting the stock bid-ask spread. On the one hand, liquidity supplying HFTs in the stock market rely on their speed advantage to better manage the adverse-selection and inventory-holding risks (see, as an example, Brogaard *et al.* 2015). On the other hand, liquidity-consuming HFTs may pick off slower traders and impose adverse-selection costs on liquidity providers (see, as an example, Shkilko & Sokolov 2020). The resulting impact on the bid-ask spread may affect the option market makers hedging costs positively or negatively depending on the underlying strategy employed by the HFT firms.

Further, the bid-ask spread may not fully capture option market makers' hedging costs as it only captures transaction costs for small orders. Lee (2008) estimates that more than half of the orders in the options markets originate from institutional investors. This, combined with the fact that option market makers typically hedge their entire inventory of options, may induce them to employ more complex execution strategies involving the use of limit and market orders and order-splitting algorithms to minimize their transaction costs. Market-making HFTs may be more willing to supply liquidity to option market-makers due to their uninformed nature and the ability of HFTs to reprice their orders continuously. On the other hand, the predictability of option market makers' hedging demand can allow liquidity-consuming HFTs to exploit the

intraday (temporary) price impact generated by option market makers hedging demand (see Van Kervel & Menkveld 2019; Yang & Zhu 2020).

In the arbitrage channel, liquidity-demanding HFTs engaging in “toxic” arbitrage strategies may expose dealers to the risk of trading at stale quotes and force them to charge a larger bid-ask spread in options markets. Given that HFTs’ liquidity demanding and supplying strategies are expected to have different effects on option spreads via the hedging and arbitrage channels, testing this hypothesis requires us to use more granular HFT data. More explicitly, we need to decompose HFT activity into liquidity-demanding and -supplying components. Unfortunately, the SEC’s MIDAS data is not granular at this level. The MIDAS supplied proxies are general measures of HFTs and include the effects of both liquidity demanding and liquidity supplying HFT activities. Therefore, in the next section, we employ a more granular dataset – the NASDAQ HFT dataset – to test these two channels.

## **5. Heterogenous HFT Strategies and Its Impact on Option Spread: Hedging and Arbitrage Channels**

### **5.1. The Description of the NASDAQ HFT Data**

The NASDAQ HFT dataset is used to analyze the two proposed channels (hedging and arbitrage) to explain the role of HFTs in the options market making. This data contains transactions for 120 randomly selected NASDAQ and NYSE-listed stocks trading during 2009. The dataset stamps transactions into those initiated by HFTs and non-HFTs.<sup>12</sup> The following variables are included: date, time (in milliseconds), trading volume, price, buy-sell indicator, and the liquidity nature of the two sides to each trade (HH, HN, NH, and NN). HH refers to a trade in which both liquidity providers and takers are HFTs. HN (NH) implies that an HFT (a non-HFT) demands liquidity, and a non-HFT (HFT) supplies liquidity. Finally, NN indicates

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<sup>12</sup> The disaggregation process is done by the NASDAQ. Brogaard *et al.* (2014) provide full details of the disaggregation.



a transaction between two non-HFTs demanding and supplying liquidity. Consistent with Brogaard *et al.* (2014), we define the sum of HH, HN, and NH as the total HFT volume. The total trading volume is about 44,800 million shares, for which 31,968 million or 71.30% have HFTs as counterparties. The total value of all trades is US\$1,381 billion.

We construct HFTs' liquidity demanding ( $SHFT_{i,d}^D$ ) and liquidity supplying ( $SHFT_{i,d}^S$ ) trades from the NASDAQ dataset.  $SHFT_{i,d}^D$  ( $SHFT_{i,d}^S$ ) is computed as the sum of HH and HN (HH and NH) divided by the total trading volume.  $SHFT_{i,d}^{All}$  is the ratio of the total HFT trading volume (HH, HN, and NH) to the total trading volume. The summary statistics of the NASDAQ HFT data is provided in Table A.5.

While the NASDAQ dataset contains transactions for 120 stocks, we only use 103 stocks as we cannot match all option contracts in OptionMetrics with the corresponding stock in the NASDAQ dataset because of inconsistencies in ticker symbols across the two datasets. Thus, we obtain the option data for these 103 stocks during 2009.

## 5.2. Estimation Approaches

Similar to the baseline model, we first estimate the OLS approach with stock and time fixed effects:

$$OSpread_{i,d} = \alpha_i + \beta_d + \gamma_1 SHFT_{i,d}^{All} + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d} \quad (4)$$

$$OSpread_{i,d} = \alpha_i + \beta_d + \gamma_1 SHFT_{i,d}^D + \gamma_2 SHFT_{i,d}^S + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d} \quad (5)$$

where  $SHFT_{i,d}^{All}$ ,  $SHFT_{i,d}^D$  and  $SHFT_{i,d}^S$  are the measures of HFTs' total, liquidity demanding, and supplying activities, respectively.  $\alpha_i$  and  $\beta_d$  are stock and time (day) fixed effects. All these variables are defined in Table 1.

As noted, to address for endogeneity, we employ 2SLS IV approach in the baseline model; the same approach is used for this analysis too:

$$OSpread_{i,d} = \alpha_i + \beta_d + \gamma_1 \widehat{SHFT}_{i,d}^{All} + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d} \quad (6)$$

$$OSpread_{i,d} = \alpha_i + \beta_d + \gamma_1 \widehat{SHFT}_{i,d}^D + \gamma_2 \widehat{SHFT}_{i,d}^S + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d} \quad (7)$$

$$SHFT_{i,d}^{All} = \alpha_i + \beta_d + \vartheta_1 IV_{i,d} + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d} \quad (8)$$

$$SHFT_{i,d}^D = \alpha_i + \beta_d + \vartheta_1 IV_{i,d} + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d} \quad (9)$$

$$SHFT_{i,d}^S = \alpha_i + \beta_d + \vartheta_1 IV_{i,d} + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d} \quad (10)$$

where  $\widehat{SHFT}_{i,d}^{All}$ ,  $\widehat{SHFT}_{i,d}^D$  and  $\widehat{SHFT}_{i,d}^S$  are the fitted values of  $SHFT_{i,d}^{All}$ ,  $SHFT_{i,d}^D$  and  $SHFT_{i,d}^S$  obtained by regressing the respective variables on  $IV_{i,d}$ . We again employ two sets of instruments for robustness. First, we use the NASDAQ HFT dataset from 2009, and any exogenous shock to the volume of HFT on the NASDAQ stock exchange during this period is a candidate for the instrument. A potential instrument satisfying these criteria is proposed by Skjeltorp *et al.* (2016). On June 5, 2009, the NASDAQ stock exchange introduced NASDAQ-Only Flash and Flash Enhanced Routable Orders.<sup>13</sup>

The implementation details for these orders and some numerical examples are also provided by Skjeltorp *et al.* (2016). Specifically, after an unsuccessful execution attempt in the NASDAQ limit order book, the NASDAQ gives an additional 500 milliseconds to its market participants and vendors to expose the orders before reaching the general marketplace. It is clear from the time constraint that only qualified low-latency traders, i.e., HFTs, are expected to use flash orders (see Harris & Namvar 2016).<sup>14</sup> This expectation is also consistent with the flash order implementation of Direct Edge – the first company to introduce flash orders on January 27, 2006 – who states that such orders allow brokers and HFTs to see and execute flash orders (see Skjeltorp *et al.* 2016). Thus, HFTs benefit from flash orders, such that the latter are expected to increase HFTs participation. In this specification,  $IV_{i,d}$  is a dummy equal *one* from June 5, 2009, to August 31, 2009, and *zero* for the other periods (from January 1, 2009, to June

<sup>13</sup> <https://www.nasdaqtrader.com/TraderNews.aspx?id=ETA2009-35>

<sup>14</sup> Some regulators and investors have argued that flash orders give an unfair advantage to market participants who are able to use them. For example, Mary Schapiro, SEC Chairman, says that “*flash orders have the potential to discourage publicly displayed trading interest and harm quote competition among markets*”.

4, 2009, and from September 1, 2009, to December 31, 2009) in our sample. It is important to note that, in this specification, we include only stock fixed effect as our instrument does not have a time variation.

Second, similar to the main analysis, inspired by Hasbrouck and Saar (2013), the level of HFT is instrumented with the average level of HFT on that day in all other stocks in the corresponding size quintile; for this specification,  $IV_{i,d}$  is the average level of three HFT proxies ( $SHFT_{i,d}^{All}$ ,  $SHFT_{i,d}^D$  and  $SHFT_{i,d}^S$ ) in all other stocks in the corresponding size quintile.

### 5.3. Estimation Results

The estimation results are presented in Table 7.

#### INSERT TABLE 7 ABOUT HERE

The first-stage results of the 2SLS IV approach are reported in the Appendix section (Table A.6). The selected instruments are significantly related to  $SHFT_{i,d}^{All}$ ,  $SHFT_{i,d}^D$  and  $SHFT_{i,d}^S$  and the signs of associations are as expected. There are three important points in Table 7. First, consistent with the results based on the SEC's MIDAS data, HFT is positively and significantly related to option spreads. The magnitude of the increase in option spreads is also economically meaningful. For example, estimates using the OLS regression with fixed effects show that a one-standard-deviation increase in  $SHFT_{i,d}^{All}$  raises  $ODspread_{i,d}$  and  $OPspread_{i,d}$  by about 14.7% and 3.0%.

. Second, consistent with the hedging channel, HFTs' liquidity-demanding orders ( $SHFT_{i,d}^D$ ) are positively and significantly (both statistically and economically) related to both option spread proxies ( $OPspread_{i,d}$  and  $ODspread_{i,d}$ ) for ATM and ITM contracts; for ITM contracts, the association between  $SHFT_{i,d}^D$  and  $OPspread_{i,d}$  is weekly significant (10% level). The economic magnitudes of the impact are again substantial. Specifically, a one-standard-deviation shock to  $SHFT_{i,d}^D$  increases  $OPspread_{i,d}$  by 5.71% and  $ODspread_{i,d}$  by 13.89% for

ATM options. For ITM options, the economic magnitudes are 6.22% for  $OPspread_{i,d}$  and 9.66% for  $ODspread_{i,d}$ . The association between  $SHFT_{i,d}^D$  and  $OSpread_{i,d}$  is not significant for OTM options. These results are consistent with the *hedging* channel.

Second, liquidity-supplying orders initiated by HFTs ( $SHFT_{i,d}^S$ ) are weakly significantly (negatively) related to both options spread proxies for ITM contracts only; it is not significant for ATM and OTM options. This result seems puzzling as it is expected that the hedging costs of the options market maker should decrease with the HFTs-liquidity supplying trading. We argue that this may be linked to the fact that we control for the stock spread in our regression setting. More explicitly, a significant proportion of the effect of  $SHFT_{i,d}^S$  on the option market spread can be captured by the bid-ask spread in the stock market. To test this argument, we estimate the models 4 to 10 without controlling for the stock spread.

#### **INSERT TABLE 8 ABOUT HERE**

Panels A ( $OPspread_{i,d}$ ) and B ( $ODspread_{i,d}$ ) of Table 8 show the estimation results for models 4 to 10 without controlling for the stock spread. The results show that both  $SHFT_{i,d}^D$  and  $SHFT_{i,d}^S$  are statistically significantly related to  $OPspread_{i,d}$  and  $ODspread_{i,d}$  in this specification. Importantly, the associations are statistically significant across all moneyness groups, albeit weakly significant in some specifications. The results are economically meaningful too. For instance, for ATM options, one standard deviation increase in  $SHFT_{i,d}^S$  ( $SHFT_{i,d}^D$ ) is associated with  $-5.00\%$  ( $+6.43\%$ ) and  $-13.89\%$  ( $+20.83\%$ ) change in  $OPspread_{i,d}$  and  $ODspread_{i,d}$ , respectively.

Overall, the results imply that HFTs' effect on the options market making is restricted to their liquidity-demanding orders in the underlying market after controlling for the stock spread. This can be explained by the fact that it is unlikely that the options market maker observes who (HFT or non-HFT) is supplying liquidity and strategically choose one over the

other, as these are anonymous markets. However, this explanation raises an interesting question about why  $SHFT_{i,d}^D$  is still significant. We argue that this is because the relationship between HFT in the stock market and the options market bid-ask spread is unlikely to be fully captured by including the stock market bid-ask spread as a control variable. More explicitly, bid-ask spreads only capture transaction costs associated with small orders and do not capture the price impact of large parent orders being split into smaller child orders.<sup>15</sup>

This is indeed a plausible explanation as Lee (2008) shows that half of the orders in the options markets originate from institutional investors who are commonly splitting their orders (see Menkveld 2008; Chemmanur *et al.* 2010). Furthermore, it is known that option market makers typically hedge their entire inventory of options, and thus, they may need to split their large orders to reduce their transaction costs. Liquidity-demanding HFTs employ different aggressive strategies (e.g., arbitrage, back-running, etc.) (see Brogaard *et al.* 2015), which allow them to profit from the price impact of options market maker's large being split into child orders. This finding further suggests that the stock spread mainly includes the impact of HFTs' liquidity supply trades on the stock market liquidity. Therefore, the studies that investigate the role of HFTs in market quality using only the spread as a proxy for liquidity should be interpreted with caution.

As noted, in addition to the hedging channel discussed and tested above, the arbitrage channel, which suggests that the option market makers can charge a higher bid-ask spread due to liquidity-consuming HFTs engaging in cross-market arbitrage strategies between stock and option markets, may also explain the impact of HFT activity in the stock market affects option market spreads. In the next analysis, we test this channel.

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<sup>15</sup> The literature examining the effect of HFT in the stock market similarly finds that HFT simultaneously leads to lower bid-ask spreads (see Hendershott *et al.* 2011; Brogaard *et al.* 2015) and higher execution costs for large orders (see Korajczyk & Murphy 2019; Van Kervel & Menkveld 2019).

### 5.3.1. Arrival of New Information and HFT

Foucault *et al.* (2017) argue that arbitrage opportunities arising due to the asynchronous adjustments in the price of correlated assets are “toxic” and lead to increased bid-ask spreads if they result from the arrival of new information.<sup>16</sup> Inspired by this, to test the arbitrage channel and investigate the relative importance of the hedging and arbitrage channels, we study the differences in the impact of HFT on the options market spread during news and no news days.

We follow Hirschey (2020) to identify the days with firm news. First, we use *Factiva*, which contains news from over 35,000 sources, and identify the news days for each firm. As suggested by Hirschey (2020), traders may respond to signals not covered by these sources. Therefore, we compute absolute market-adjusted returns in a second step and exclude days on which they were greater than specific thresholds (1%, 0.5%, and 0.25%).<sup>17</sup>

#### INSERT TABLE 9 ABOUT HERE

Panels A and B of Table 9 present the model’s estimation results for days with no firm news. The results show that  $SHFT_{i,d}^D$  has a positive and statistically significant influence on the ATM options spread, even on these days. Thus, the hedging channel is more dominant for ATM options as option market makers need to rebalance their entire inventory positions. In contrast, the relationship between  $SHFT_{i,d}^D$  and the ITM options spread loses its significance once we exclude days with firm news. Halpern and Turnbull (1985) and Galai (1978) document that the profitable arbitrage opportunities arise from violations put-call parity are more frequently observed in ITM options. Therefore, the association between HFT and option spreads disappears on days with fewer arbitrage opportunities (i.e., without firm news) suggests that the *arbitrage* channel is more dominant for ITM contracts.

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<sup>16</sup> This is also consistent with Rzayev and Ibikunle (2019) and Brolley and Zoican (2020), who show that HFTs can make profits at slow traders’ expense due to their ability to react faster to public news (i.e., *latency arbitrage*).

<sup>17</sup> Our sample period covers 2009. Due to the 2008 Financial Crisis, there may be some abnormal returns during our sample period not related to any specific information. Therefore, for robustness, in the second specification, we exclude days with market-adjusted returns higher than the mean of the market-adjusted returns for the sample period. The results obtained are qualitatively similar to those we report.

It is important to note that we do not claim that the arbitrage channel is the only channel that can explain the association between HFT and ITM option spreads. This is because the NASDAQ HFT data contains a limited number of stocks, and it covers one year – 2009. Moreover, the data includes HFT activity in NASDAQ only. In the main analysis, where we use the comprehensive panel data from the SEC’s MIDAS database, we find that the economic magnitude of the effect of HFT on option spreads is quite substantial for ITM options. Thus, it is plausible to expect that both hedging and arbitrage channels contribute to the association between HFT and the costs of market making in ITM options. We nevertheless believe that the fact that the relationship between  $SHFT_{i,d}^D$  and the ITM option spread loses its significance in no news days gives us some insights about the association between HFT activity and ITM option spreads. Thus, it should be of general interest.

## 6. Conclusion

This study uses a comprehensive sample of HFT data provided by the SEC MIDAS to investigate the relation between HFT activity in the stock market and its impact on options market making. We find that HFT in the stock market leads to increased bid-ask spread and deterioration of liquidity in options markets. We propose two channels to explain this finding: (i) hedging channel and (ii) arbitrage channel. The hedging channel suggests that the options market maker hedging costs increase (decrease) due to higher (lower) bid-ask spread and price impact sourced HFTs’ liquidity demanding (supplying) trades. On the other hand, the arbitrage channel implies that HFTs’ aggressive (liquidity demanding) trades expose the options market maker to the risk of trading at stale prices upon the arrival of new information.

We test these channels by using proprietary HFT data from NASDAQ. Our findings suggest that the hedging channel dominates for ATM options. On the other hand, the significant relation between HFT and the options spread might be mainly driven by the arbitrage channel

for ITM options. For OTM contracts, the economic magnitude of the impact of HFT on option spreads is lower in the main analysis and is not statistically significant based on the NASDAQ HFT data.

Our findings also suggest that while HFTs' liquidity-supplying trades increase the options market maker's hedging abilities by improving liquidity in the underlying markets, the significant relation between HFTs' liquidity supply trades and options market liquidity disappears after controlling for the stock spread.

The results of this paper highlight the necessity for a better understanding of the costs/risks due to HFTs in today's highly fragmented and complicated market structures. Specifically, our findings suggest that practitioners, academics, and policy makers should carefully consider the cross-asset effects of HFTs activities in equity markets on derivatives market quality.



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**Table 1. Definitions of variables**

This table reports the notation, description, and source of variables. The units of the variables are in parentheses following the variable names. Panel A reports the equity market variables and Panel B reports options market variables.

Panel A. Equity market variables		
Variable	Description	Data source
$QT_{i,d}$	Quote to trade ratio for firm $i$ in day $d$ is computed as the sum of order volume for all order messages divided by the sum of trade volume for all trades that are not against hidden orders.	MIDAS
$CT_{i,d}$ (%)	Cancel to trade for firm $i$ in day $d$ is computed as the number of all cancel messages (full or partial) divided by the number of all trade messages.	MIDAS
$OR_{i,d}$ (%)	Odd lot rate for firm $i$ in day $d$ is computed as the number of odd lot trade messages for all exchanges divided by the number of trades from exchange that report individual trades.	MIDAS
$OV_{i,d}$ (%)	Odd lot rate for firm $i$ in day $d$ is computed as the sum of odd lot trade volume for all exchanges divided by the sum of trade volume from exchange that report individual trades.	MIDAS
$ITS_{i,d}$	Inverse trade size for firm $i$ in day $d$ is computed as the number of trades divided by the number of shares traded.	MIDAS
$SHFT_{i,d}^{All}$ (%)	HFTs' total trades percentage for stock $i$ and day $d$ computed as the ratio of the HFTs' total trading volume (the sum of HH, HN and NH) to the total trading volume.	NASDAQ
$SHFT_{i,d}^D$ (%)	HFTs' liquidity-demanding trades percentage for stock $i$ and day $d$ computed as the ratio of the HFTs' liquidity-demanding trading volume (the sum of HH and HN) to the total trading volume.	NASDAQ
$SHFT_{i,d}^S$ (%)	HFTs' liquidity-supplying trades percentage for stock $i$ and day $d$ computed as the ratio of the HFTs' liquidity-supplying trading volume (the sum of HH and NH) to the total trading volume.	NASDAQ
$SDspread_{i,d}$	Stock dollar spread for firm $i$ in day $d$ is computed as the difference between best ask and bid prices for day $d$ .	CRSP
$SPspread_{i,d}$ (%)	Stock proportional spread for firm $i$ in day $d$ is computed as bid-ask spread divided by bid-ask midpoint for stock $i$ in each day.	CRSP
$SVolatility_{i,d}$	Stock price volatility for firm $i$ in day $t$ is computed as the absolute value of the difference between mid-prices for days $d$ and $d-1$ .	CRSP
Panel B. Option market variables		
$ODspread_{i,d}$	Option market dollar spread for stock $i$ and day $d$ computed as the dollar-volume-weighted average of the dollar spread (the difference between the best ask and bid prices).	OptionMetrics
$OPspread_{i,d}$ (%)	Option market proportional spread for stock $i$ and day $d$ computed as the dollar-volume-weighted average of the proportional spread (the difference between the best ask and bid prices divided by the midpoint of the ask and bid prices).	OptionMetrics
$Ovolume_{i,d}$	Option volume for stock $i$ and day $d$ computed as the natural logarithm of the daily trading volume (contracts)	OptionMetrics
$Oimplied_{i,d}$	Option-implied volatility for stock $i$ and day $d$ computed as the dollar-volume-weighted average of the implied volatility provided by OptionMetrics.	OptionMetrics
$ Odelta_{i,d} $	Absolute option delta for stock $i$ and day $d$ computed as the dollar-volume-weighted average of the absolute delta provided by OptionMetrics.	OptionMetrics
$Ogamma_{i,d}$	Option gamma for stock $i$ and day $d$ computed as the dollar-volume-weighted average of the gamma provided by OptionMetrics.	OptionMetrics
$Ovega_{i,d}$	Option vega for stock $i$ and day $d$ computed as the dollar-volume-weighted average of the vega provided by OptionMetrics.	OptionMetrics

**Table 2. Summary statistics for MIDAS Sample**

This table reports the descriptive statistics for the variables used in our analysis. Panel A shows the descriptive statistics for all variables from the underlying stock market. Panel B provides the descriptive statistics for all options-related variables separately for the full sample and three groups based on moneyness. For the definitions and computation methods of the variables, see Table 1. We follow Bollen and Whaley (2004) and define OTM options as those with absolute option delta  $|\Delta| \leq 0.375$ , ATM options as those with  $0.375 < |\Delta| \leq 0.625$ , and ITM options as those with  $|\Delta| > 0.625$ . In both panels, we have two samples: (i) the full sample and (ii) the tick size pilot sample. The full sample contains 2,746 stocks traded between January 1, 2012, and December 31, 2019, on the US exchanges. The tick size pilot sample contains 1,235 stocks (617 control stocks and 618 treated stocks) included the Tick Size Pilot Programme launched by the SEC. The Tick Size Pilot Programme commenced on October 3, 2016 and implemented it over the next two years. Therefore, the tick size pilot sample covers from October 1, 2014, to September 28, 2018.

Panel A. Equity market variables				
	Variable	Mean	Median	Stdev
Full Sample	$QT_{i,d}$	42.07	33.73	29.43
	$CT_{i,t}$ (%)	23.15	19.69	13.51
	$OR_{i,t}$ (%)	37.21	35.07	17.44
	$OV_{i,t}$ (%)	15.87	14.09	9.81
	$ITS_{i,t}$	0.01	0.01	0.01
	$SDspread_{i,t}$	0.02	0.01	0.03
	$SPspread_{i,t}$ (%)	0.11	0.05	0.18
	$SVolatility_{i,t}$	0.69	0.35	1.02
Tick Size Pilot Sample	$QT_{i,d}$	44.04	32.24	36.81
	$CT_{i,t}$ (%)	24.18	19.68	16.15
	$OR_{i,t}$ (%)	39.15	38.33	14.59
	$OV_{i,t}$ (%)	16.83	15.96	8.35
	$ITS_{i,t}$	0.01	0.01	0.01
	$SDspread_{i,t}$	0.03	0.02	0.03
	$SPspread_{i,t}$ (%)	0.18	0.09	0.26
	$SVolatility_{i,t}$	0.60	0.32	0.81
Panel B. Option market variables				
Full sample	$ODspread_{i,t}$	0.45	0.27	0.55
	$OPspread_{i,t}$ (%)	34.07	19.63	40.75
	$Ovolume_{i,t}$	5.06	4.96	2.59
	$Oimplied_{i,t}$	0.45	0.36	0.28
	$ Odelta_{i,t} $	0.51	0.50	0.17
	$Ogamma_{i,t}$	0.12	0.08	0.12
	$Ovega_{i,t}$	5.98	4.10	6.25
Tick Size Pilot Sample	$ODspread_{i,t}$	0.60	0.40	0.62
	$OPspread_{i,t}$ (%)	50.44	32.43	47.79
	$Ovolume_{i,t}$	3.62	3.58	1.95
	$Oimplied_{i,t}$	0.52	0.43	0.28
	$ Odelta_{i,t} $	0.50	0.49	0.19
	$Ogamma_{i,t}$	0.12	0.10	0.10
	$Ovega_{i,t}$	4.06	2.86	3.79
	$ODspread_{i,t}$	0.38	0.23	0.47
	$OPspread_{i,t}$ (%)	21.64	13.57	25.30

ATM	$Ovolume_{i,t}$	4.53	4.54	2.42
	$Oimplied_{i,t}$	0.40	0.33	0.23
	$ Odelta_{i,t} $	0.49	0.49	0.05
	$Ogamma_{i,t}$	0.13	0.09	0.12
	$Ovega_{i,t}$	8.23	5.79	8.29
ITM	$ODspread_{i,t}$	0.62	0.39	0.71
	$OPspread_{i,t}(\%)$	15.73	10.57	16.38
	$Ovolume_{i,t}$	3.81	3.64	2.31
	$Oimplied_{i,t}$	0.48	0.38	0.31
	$ Odelta_{i,t} $	0.78	0.78	0.08
	$Ogamma_{i,t}$	0.11	0.07	0.11
	$Ovega_{i,t}$	4.86	3.11	0.11
OTM	$ODspread_{i,t}$	0.27	0.17	0.32
	$OPspread_{i,t}(\%)$	52.10	32.50	52.08
	$Ovolume_{i,t}$	4.34	4.17	2.50
	$Oimplied_{i,t}$	0.43	0.35	0.26
	$ Odelta_{i,t} $	0.25	0.26	0.06
	$Ogamma_{i,t}$	0.10	0.06	0.10
	$Ovega_{i,t}$	5.93	4.13	6.05



**Table 3. The relationship between control variables and option spread – MIDAS Sample**

This table presents the results for the estimation of the association between control variables and the options spread:

$$OSpread_{i,d} = \alpha_i + \beta_d + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d}$$

where  $OSpread_{i,d}$  corresponds to either the proportional spread ( $OPspread_{i,d}$ ) or the dollar spread ( $ODspread_{i,d}$ ).  $\alpha_i$  and  $\beta_d$  are stock and time (day) fixed effects. The  $C_{k,i,d}$  is a set of  $k$  control variables, including variables from both the option and underlying markets. The option market variables are  $Ovolume_{i,d}$ ,  $Oimplied_{i,d}$ ,  $|Odelta_{i,d}|$ ,  $Ogamme_{i,d}$  and  $Ovega_{i,d}$  and the stock market variables are  $SPspread_{i,d}$  (when we employ  $OPspread_{i,d}$  as the dependent variable),  $SDspread_{i,d}$  (when we use  $ODspread_{i,d}$  as the dependent variable), and  $SVolatility_{i,d}$ . For the definitions and computation methods of all the variables, see Table 1. The sample contains 2,746 stocks traded between January 1, 2012, and December 31, 2019, on the US exchanges. Standard errors are double clustered on stock and day and t-statistics are reported in parentheses. \*, \*\* and \*\*\* denote significance at 10%, 5% and 1%.

	$ODspread_{i,t}$	$OPspread_{i,t}$
$Ovolume_{i,d}$	-0.24*** (-256.37)	-25.55*** (-351.10)
$Oimplied_{i,d}$	0.23*** (211.08)	13.44*** (187.84)
$ Odelta_{i,d} $	0.42*** (455.23)	21.18*** (316.27)
$Ogamme_{i,d}$	-0.09*** (-224.78)	9.56*** (170.34)
$Ovega_{i,d}$	0.16*** (191.56)	-13.72*** (301.59)
$SPspread_{i,d}$		2.38*** (51.55)
$SDspread_{i,d}$	0.05*** (78.00)	
$SVolatility_{i,d}$	0.06*** (132.43)	3.89*** (170.27)
Time and Stock FEs	Yes	Yes
N	2,969,829	2,969,829

**Table 4. The impact of HFT on option spread – MIDAS Sample**

This table presents the results for the estimation of the impact of HFT on the options spread:

$$OSpread_{i,d} = \alpha_i + \beta_d + \gamma_1 \widehat{HFT}_{i,d} + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d}$$

$$HFT_{i,d} = \alpha_i + \beta_d + \vartheta_1 IV_{i,d} + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d}$$

where  $OSpread_{i,d}$  corresponds to either the proportional spread ( $OPspread_{i,d}$ ) or the dollar spread ( $ODspread_{i,d}$ ),  $HFT_{i,d}$  corresponds one of the five HFT proxies ( $QT_{i,d}$ ,  $CT_{i,d}$ ,  $OR_{i,d}$ ,  $OV_{i,d}$ ,  $ITS_{i,d}$ ).  $\alpha_i$  and  $\beta_d$  are stock and time (day) fixed effects. The  $C_{k,i,d}$  is a set of  $k$  control variables, including variables from both the option and underlying markets. The option market variables are  $Ovolume_{i,d}$ ,  $Oimplied_{i,d}$ ,  $|Odelta_{i,d}|$ ,  $Ogamme_{i,d}$  and  $Ovega_{i,d}$  and the stock market variables are  $SPspread_{i,d}$  (when we employ  $OPspread_{i,d}$  as the dependent variable),  $SDspread_{i,d}$  (when we use  $ODspread_{i,d}$  as the dependent variable), and  $SVolatility_{i,d}$ . For the definitions and computation methods of all the variables, see Table 1. Three specifications of the model are estimated. In Column 1, we estimate the model by using OLS with stock and fixed effects. In Columns 2 and 3, we use 2SLS IV approach. In Column 2, the level of HFT is instrumented with the changes in tick size; for this specification,  $IV_{i,d}$  is a dummy variable that takes the value of one after the change (from October 3, 2016 to September 28, 2018), and zero before (from October 1, 2014 to October 2, 2016) for the treatment stocks and  $IV_{i,d}$  takes the value of zero in the entire period for the control stocks. In Column 3, the level of HFT is instrumented with the average level of HFT on that day in all other stocks in the corresponding size quintile; for this specification,  $IV_{i,d}$  is the average level of five HFT proxies ( $QT_{i,d}$ ,  $CT_{i,d}$ ,  $OR_{i,d}$ ,  $OV_{i,d}$ ,  $ITS_{i,d}$ ) in all other stocks in the corresponding size quintile. For Columns 1 and 3, the sample contains 2,746 stocks traded between January 1, 2012, and December 31, 2019, on the US exchanges. For Column 2, the sample contains 1,235 stocks (617 control stocks and 618 treated stocks) that implemented in the SEC's Tick Size Pilot Programme from October 1, 2014, to September 28, 2018. Standard errors are double clustered on stock and day and t-statistics are reported in parentheses. \*, \*\* and \*\*\* denote significance at 10%, 5% and 1%.

Panel A: $ODspread_{i,t}$ is the dependent variable.				
	Variable	OLS (1)	IV (Tick Size Pilot) (2)	IV (Average HFT) (3)
$HFT_{i,d}$	$QT_{i,d}$	0.02*** (28.42)	0.02*** (10.95)	0.03*** (25.16)
	$CT_{i,d}$	0.01*** (15.52)	0.02*** (11.12)	0.01*** (10.10)
	$OR_{i,d}$	0.07*** (60.39)	0.02*** (6.85)	0.11*** (65.74)
	$OV_{i,d}$	0.05*** (54.45)	0.02*** (8.52)	0.09*** (62.46)
	$ITS_{i,d}$	0.05*** (41.60)	0.02*** (6.30)	0.08*** (44.01)
	Controls	Yes	Yes	Yes
	Time and Stock FEs	Yes	Yes	Yes
	N	2,969,829	640,306	2,967,095
Panel B: $OPspread_{i,t}$ is the dependent variable.				
	Variable	OLS (1)	IV (Tick Size Pilot) (2)	IV (Average HFT) (3)
$HFT_{i,d}$	$QT_{i,d}$	0.89*** (19.51)	1.46*** (11.19)	1.31*** (17.98)
	$CT_{i,d}$	0.60*** (11.99)	1.35*** (11.59)	0.96*** (11.65)
	$OR_{i,d}$	2.76*** (32.88)	1.67*** (5.81)	3.59*** (28.06)
	$OV_{i,d}$	1.89*** (27.72)	1.50*** (7.17)	2.61*** (25.39)
	$ITS_{i,d}$	1.13*** (12.45)	1.54*** (5.18)	2.53*** (19.84)
	Controls	Yes	Yes	Yes
	Time and Stock FEs	Yes	Yes	Yes
	N	2,969,829	640,306	2,967,095



Time and Stock FEs	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	
N	2,266,764	402,656	2,264,040	2,080,605	343,460	2,077,881	2,543,347	480,324	2,540,617	
Panel B: $OPspread_{i,t}$ is the dependent variable.										
		ATM			ITM			OTM		
$HFT_{i,t}$	OLS (1)	IV (Tick Size Pilot) (2)	IV (Average HFT) (3)	OLS (4)	IV (Tick Size Pilot) (5)	IV (Average HFT) (6)	OLS (7)	IV (Tick Size Pilot) (8)	IV (Average HFT) (9)	
$QT_{i,t}$	1.10*** (30.47)	1.24*** (9.38)	1.82*** (32.13)	0.75*** (30.74)	1.30*** (10.27)	1.19*** (31.19)	1.28*** (22.80)	2.35*** (11.19)	2.28*** (25.60)	
$CT_{i,t}$	1.21*** (30.53)	0.97*** (9.06)	2.04*** (31.64)	0.64*** (24.00)	0.62** (7.63)	1.03*** (23.90)	1.63*** (26.55)	2.24*** (11.90)	3.12*** (31.02)	
$OR_{i,t}$	2.61*** (38.34)	1.28*** (3.29)	4.00*** (39.78)	1.78*** (40.27)	1.85*** (4.79)	3.02*** (46.83)	0.92*** (8.80)	1.66*** (5.42)	2.05*** (13.31)	
$OV_{i,t}$	1.70*** (30.84)	1.46*** (5.34)	2.80*** (34.67)	1.21*** (33.83)	2.35*** (8.06)	2.12*** (41.30)	-0.09 (-1.02)	1.57*** (5.51)	0.66*** (5.35)	
$ITS_{i,t}$	1.55*** (20.71)	1.09*** (2.67)	3.10*** (30.46)	0.97*** (19.45)	4.44*** (7.78)	1.99*** (29.69)	-1.36*** (-12.06)	1.72*** (4.17)	-0.19 (-1.25)	
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	
Time and Stock FEs	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	
N	2,266,764	402,656	2,264,040	2,080,605	343,460	2,077,881	2,543,347	480,324	2,540,617	

**Table 6. Economic effect of HFT on option spreads – MIDAS Sample**

This table presents the economic effect of one standard deviation shock to HFT proxies on the option spreads:

$$\text{Economic effect} = \gamma_1 / \mu(\text{Ospread}_{i,t}),$$

where  $\gamma_1$  is the coefficient of various HFT proxies obtain from estimation of the regression,  $\mu(\text{Ospread}_{i,t})$  is the average value of various  $\text{ODspread}_{i,t}$  and  $\text{OPspread}_{i,t}$ . In Panel A (B) the economic magnitude of the impact of HFT on the  $\text{ODspread}_{i,t}$  ( $\text{OPspread}_{i,t}$ ). The first column shows the results for standard OLS approach with stock and day fixed effects, whereas the second and third columns present the results for 2SLS IV approach. In Column 2, the level of HFT is instrumented with the changes in tick size; for this specification,  $\text{IV}_{i,d}$  is a dummy variable that takes the value of one after the change (from October 3, 2016 to September 28, 2018), and zero before (from October 1, 2014 to October 2, 2016) for the treatment stocks and  $\text{IV}_{i,d}$  takes the value of zero in the entire period for the control stocks. In Column 3, the level of HFT is instrumented with the average level of HFT on that day in all other stocks in the corresponding size quintile; for this specification,  $\text{IV}_{i,d}$  is the average level of five HFT proxies ( $\text{QT}_{i,d}$ ,  $\text{CT}_{i,d}$ ,  $\text{OR}_{i,d}$ ,  $\text{OV}_{i,d}$ ,  $\text{ITS}_{i,d}$ ) in all other stocks in the corresponding size quintile. For Columns 1 and 3, the sample contains 2,746 stocks traded between January 1, 2012, and December 31, 2019, on the US exchanges. For Column 2, the sample contains 1,235 stocks (617 control stocks and 618 treated stocks) that implemented in the SEC's Tick Size Pilot Programme from October 1, 2014, to September 28, 2018. We follow Bollen and Whaley (2004) and define OTM options as those with  $|\text{Odelta}_{i,t}| \leq 0.375$ , ATM options as those with  $0.375 < |\text{Odelta}_{i,t}| \leq 0.625$ , and ITM options as those with  $|\text{Odelta}_{i,t}| > 0.625$ . For the definitions and computation methods of all the variables, see Table 1.

Panel A: $\text{ODspread}_{i,t}$ is the dependent variable.												
	OLS (1)				IV (Tick Size Pilot) (2)				IV (Average HFT) (3)			
$\text{HFT}_{i,t}$	Full	ATM	ITM	OTM	Full	ATM	ITM	OTM	Full	ATM	ITM	OTM
$\text{QT}_{i,t}$	4.44%	5.26%	6.45%	3.70%	3.33%	3.64%	5.81%	2.63%	6.67%	5.26%	9.68%	7.41%
$\text{CT}_{i,t}$	2.22%	2.63%	1.61%	3.70%	3.33%	3.64%	2.33%	2.63%	2.22%	5.26%	1.61%	3.70%
$\text{OR}_{i,t}$	15.56%	10.53%	29.03%	7.41%	3.33%	5.45%	9.30%	2.63%	24.44%	18.42%	43.55%	11.11%
$\text{OV}_{i,t}$	11.11%	7.89%	24.19%	3.70%	3.33%	5.45%	10.47%	2.63%	20.00%	13.16%	37.10%	7.41%
$\text{ITS}_{i,t}$	11.11%	5.26%	24.19%	1.11%	3.33%	5.45%	17.44%	2.63%	17.78%	10.53%	35.48%	3.70%
Average	8.89%	6.32%	17.10%	3.93%	3.33%	4.73%	9.07%	2.63%	14.22%	10.53%	25.48%	6.67%
Panel B: $\text{OPspread}_{i,t}$ is the dependent variable.												
	OLS (1)				IV (Tick Size Pilot) (2)				IV (Average HFT) (3)			
$\text{HFT}_{i,t}$	Full	ATM	ITM	OTM	Full	ATM	ITM	OTM	Full	ATM	ITM	OTM
$\text{QT}_{i,t}$	2.61%	5.08%	4.77%	2.46%	2.89%	3.47%	5.16%	3.04%	3.85%	8.41%	7.57%	4.38%
$\text{CT}_{i,t}$	1.76%	5.59%	4.07%	3.13%	2.68%	2.71%	2.46%	1.60%	2.82%	9.43%	6.55%	5.99%
$\text{OR}_{i,t}$	8.10%	12.06%	11.32%	1.77%	3.31%	3.58%	7.35%	2.15%	10.54%	18.48%	19.20%	3.93%
$\text{OV}_{i,t}$	5.55%	7.86%	7.69%	-0.17%	2.97%	4.08%	9.33%	2.03%	7.66%	12.94%	13.48%	1.27%
$\text{ITS}_{i,t}$	3.32%	7.16%	6.17%	-2.61%	3.05%	3.05%	17.63%	2.22%	7.43%	14.33%	12.65%	-0.36%
Average	4.27%	7.55%	6.80%	0.91%	2.98%	3.38%	8.39%	2.21%	6.46%	12.72%	11.89%	3.04%

**Table 7. The impact of HFT on option spread – NASDAQ Sample**

This table presents the results for the estimation of the impact of HFT on the options spread:

$$OSpread_{i,d} = \alpha_i + \beta_d + \gamma_1 \widehat{SHFT}_{i,d}^D + \gamma_2 \widehat{SHFT}_{i,d}^S + \sum_{k=1}^6 \delta_k C_{k,i,d} + \varepsilon_{i,d}$$

$$SHFT_{i,d}^D = \alpha_i + \beta_d + \vartheta_1 IV_{i,d} + \sum_{k=1}^6 \delta_k C_{k,i,d} + \varepsilon_{i,d}$$

$$SHFT_{i,d}^S = \alpha_i + \beta_d + \vartheta_1 IV_{i,d} + \sum_{k=1}^6 \delta_k C_{k,i,d} + \varepsilon_{i,d}$$

where  $OSpread_{i,d}$  corresponds to either the proportional spread ( $OPspread_{i,d}$ ) or the dollar spread ( $ODspread_{i,d}$ ),  $SHFT_{i,d}^D$  and  $SHFT_{i,d}^S$  are the measures of HFTs-liquidity demanding and supplying activities, respectively.  $\alpha_i$  and  $\beta_d$  are stock and time (day) fixed effects. The  $C_{k,i,d}$  is a set of  $k$  control variables, including variables from both the option and underlying markets. The option market variables are  $Ovolume_{i,d}$ ,  $Oimplied_{i,d}$ ,  $|Odelta_{i,d}|$ ,  $Ogamme_{i,d}$  and  $Ovega_{i,d}$  and the stock market variables are  $SPspread_{i,d}$  (when we employ  $OPspread_{i,d}$  as the dependent variable),  $SDspread_{i,d}$  (when we use  $ODspread_{i,d}$  as the dependent variable), and  $SVolatility_{i,d}$ . For the definitions and computation methods of all the variables, see Table 1. Three specifications of the model are estimated. In Columns 1 and 4, we estimate the model by using OLS with stock and fixed effects. In Columns 2 and 5, we use 2SLS IV approach.  $IV_{i,d}$  is a dummy variable that takes the value 1 during the flash-orders period (from June 5, 2009, to August 31, 2009) initiated by the NASDAQ. In Columns 3 and 6, the level of HFT is instrumented with the average level of HFT on that day in all other stocks in the corresponding size quintile; for this specification,  $IV_{i,d}$  is the average level of two HFT proxies ( $SHFT_{i,d}^D$  and  $SHFT_{i,d}^S$ ) in all other stocks in the corresponding size quintile. The sample contains 103 stocks traded between January 1, 2009, and December 31, 2009, on the NASDAQ. We follow Bollen and Whaley (2004) and define OTM options as those with  $|Odelta_{i,t}| \leq 0.375$ , ATM options as those with  $0.375 < |Odelta_{i,t}| \leq 0.625$ , and ITM options as those with  $|Odelta_{i,t}| > 0.625$ . Standard errors are double clustered on stock and day, and t-statistics are reported in parentheses. \*, \*\* and \*\*\* denote significance at 10%, 5% and 1%.

Panel A: $ODspread_{i,t}$ is the dependent variable.									
	$SHFT_{i,d}^{All}$			$SHFT_{i,d}^D$			$SHFT_{i,d}^S$		
	OLS (1)	IV (Flash orders) (2)	IV (Average HFT) (3)	OLS (1)	IV (Flash orders) (2)	IV (Average HFT) (3)	OLS (1)	IV (Flash orders) (2)	IV (Average HFT) (3)
Full	0.02*** (3.93)	0.02*** (3.85)	0.02*** (3.41)	0.03*** (3.11)	0.03*** (3.27)	0.03*** (3.57)	-0.01 (-1.26)	-0.01 (-1.33)	-0.01 (-1.28)
ATM	0.02** (1.98)	0.02** (2.17)	0.02** (2.22)	0.02** (2.28)	0.03*** (2.80)	0.03*** (2.89)	-0.01 (-1.05)	-0.01 (-1.17)	-0.01 (-1.06)
ITM	0.01* (1.80)	0.01* (1.83)	0.01* (1.82)	0.02*** (2.58)	0.02*** (2.63)	0.02*** (2.91)	-0.01** (-2.38)	-0.01* (-1.77)	-0.01* (-1.85)
OTM	0.01 (0.33)	0.00 (0.04)	0.01 (0.16)	0.01 (0.94)	0.01 (0.55)	0.01 (1.12)	-0.01 (-0.77)	-0.01 (-0.90)	-0.01 (-0.83)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Stock FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Time FE	Yes	No	Yes	Yes	No	Yes	Yes	No	Yes







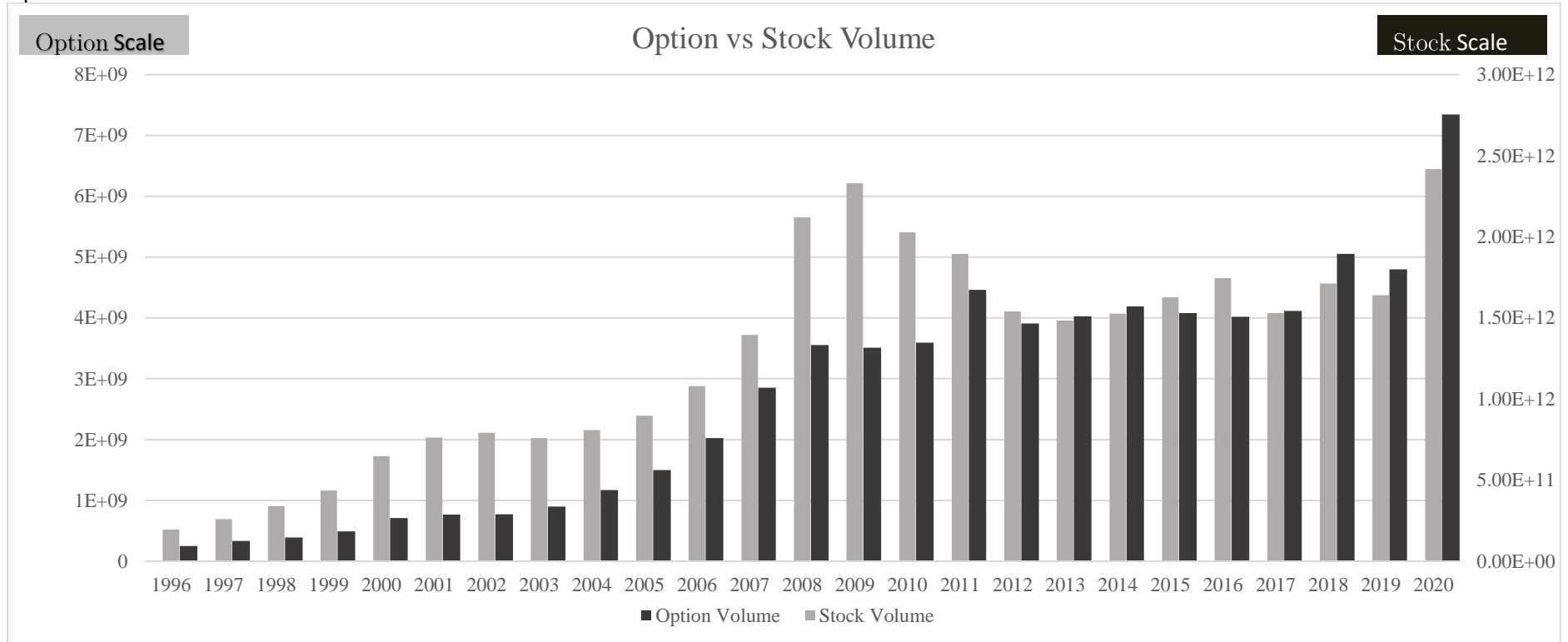






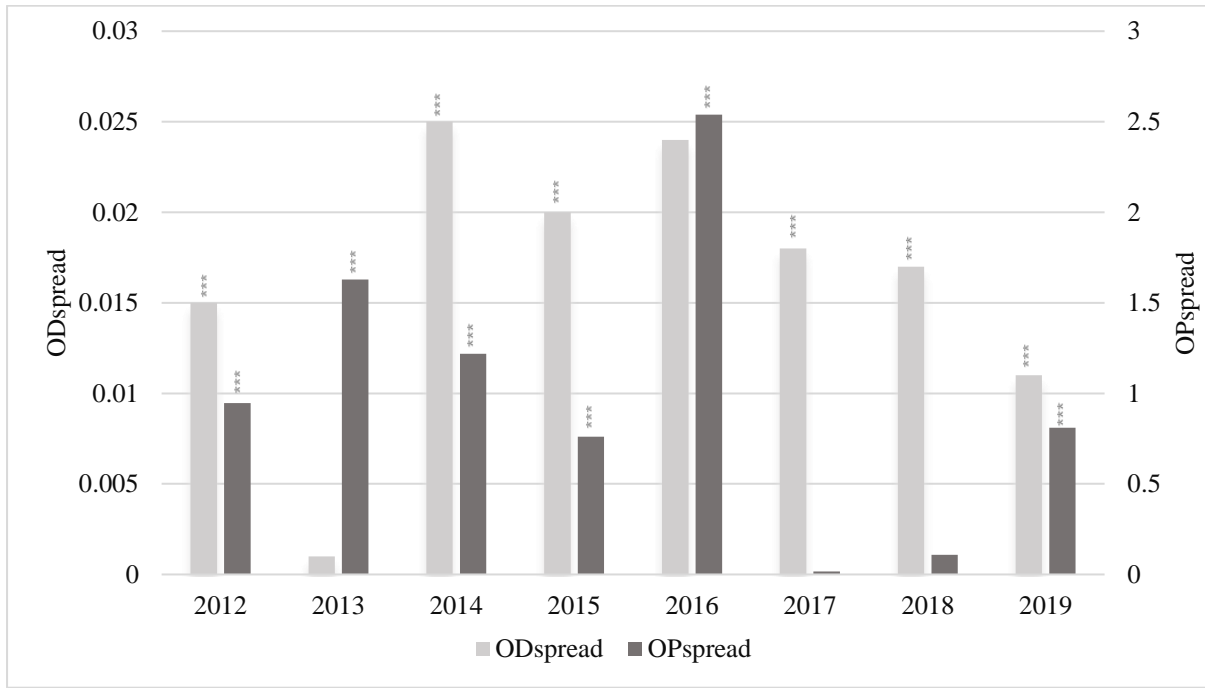
**Figure 1. The evolution of trading volume in US equity and options markets**

This figure reports the evolution of trading volume in US equity and options markets. The grey (dark) bar corresponds to the number of shares (contracts) traded in US equity (options) markets. The sample contains all stocks traded between January 1, 1996, and December 31, 2020, on the US exchanges. The data is obtained from CRSP and Optionmetrics.



**Figure 2. The impact of HFT on option spread – MIDAS Sample: Year by year estimation**

This figure reports the results from year-by-year OLS regressions of the association between HFT and the options spread. The grey (dark) bar corresponds to the impact of HFT on the options dollar (proportional) spread.  $QT_{i,d}$  is used as an HFT measure ( $\gamma_1$ ). The sample contains 2,746 stocks traded between January 1, 2012, and December 31, 2019, on the US exchanges. All estimations include stock and time (day) fixed effects. Standard errors are double clustered on stock and day and t-statistics are reported in parentheses. \*, \*\* and \*\*\* denote significance at 10%, 5% and 1%.



**Table A.1. The impact of HFT on option spread controlling for the volatility of stock spread – MIDAS Sample**

This table presents the results for the estimation of the impact of HFT on the options spread:

$$OSpread_{i,d} = \alpha_i + \beta_d + \gamma_1 HFT_{i,d} + \sum_{k=1}^8 \delta_k C_{k,i,d} + \varepsilon_{i,d}$$

$$HFT_{i,d} = \alpha_i + \beta_d + \vartheta_1 IV_{i,d} + \sum_{k=1}^8 \delta_k C_{k,i,d} + \varepsilon_{i,d}$$

where  $OSpread_{i,d}$  corresponds to either the proportional spread ( $OPspread_{i,d}$ ) or the dollar spread ( $ODspread_{i,d}$ ),  $HFT_{i,d}$  corresponds one of the five HFT proxies ( $QT_{i,d}$ ,  $CT_{i,d}$ ,  $OR_{i,d}$ ,  $OV_{i,d}$ ,  $ITS_{i,d}$ ).  $\alpha_i$  and  $\beta_d$  are stock and time (day) fixed effects. The  $C_{k,i,d}$  is a set of  $k$  control variables, including variables from both the option and underlying markets. The option market variables are  $Ovolume_{i,d}$ ,  $Oimplied_{i,d}$ ,  $|Odelta_{i,d}|$ ,  $Ogamme_{i,d}$  and  $Ovega_{i,d}$  and the stock market variables are  $SPspread_{i,d}$  and  $SPspreadvol_{i,d}$  (when we employ  $OPspread_{i,d}$  as the dependent variable),  $SDspread_{i,d}$  and  $SDspreadvol_{i,d}$  (when we use  $ODspread_{i,d}$  as the dependent variable), and  $SVolatility_{i,d}$ .  $SPspreadvol_{i,d}$  and  $SDspreadvol_{i,d}$  are the volatility of proportional and dollar spreads. For the definitions and computation methods of all the other variables, see Table 1. Three specifications of the model are estimated. In Column 1, we estimate the model by using OLS with stock and fixed effects. In Columns 2 and 3, we use 2SLS IV approach. In Column 2, the level of HFT is instrumented with the changes in tick size; for this specification,  $IV_{i,d}$  is a dummy variable that takes the value of one after the change (from October 3, 2016 to September 28, 2018), and zero before (from October 1, 2014 to October 2, 2016) for the treatment stocks and  $IV_{i,d}$  takes the value of zero in the entire period for the control stocks. In Column 3, the level of HFT is instrumented with the average level of HFT on that day in all other stocks in the corresponding size quintile; for this specification,  $IV_{i,d}$  is the average level of five HFT proxies ( $QT_{i,d}$ ,  $CT_{i,d}$ ,  $OR_{i,d}$ ,  $OV_{i,d}$ ,  $ITS_{i,d}$ ) in all other stocks in the corresponding size quintile. For Columns 1 and 3, the sample contains 2,746 stocks traded between January 1, 2012, and December 31, 2019, on the US exchanges. For Column 2, the sample contains 1,235 stocks (617 control stocks and 618 treated stocks) that implemented in the SEC's Tick Size Pilot Programme from October 1, 2014, to September 28, 2018. Standard errors are double clustered on stock and day and t-statistics are reported in parentheses. \*, \*\* and \*\*\* denote significance at 10%, 5% and 1%.

Panel A: $ODspread_{i,t}$ is the dependent variable.				
	Variable	OLS (1)	IV (Tick Size Pilot) (2)	IV (Average HFT) (3)
$HFT_{i,d}$	$QT_{i,d}$	0.02*** (26.51)	0.02*** (9.81)	0.03*** (25.99)
	$CT_{i,d}$	0.01*** (13.72)	0.02*** (10.01)	0.01*** (9.25)
	$OR_{i,d}$	0.07*** (57.68)	0.02*** (6.71)	0.11*** (64.80)
	$OV_{i,d}$	0.05*** (53.38)	0.02*** (8.49)	0.09*** (62.55)
	$ITS_{i,d}$	0.05*** (41.03)	0.02*** (6.22)	0.08*** (44.36)
	Controls	Yes	Yes	Yes
Time and Stock FEs		Yes	Yes	Yes
N		2,969,829	640,306	2,967,095
Panel B: $OPspread_{i,t}$ is the dependent variable.				
	Variable	OLS (1)	IV (Tick Size Pilot) (2)	IV (Average HFT) (3)
$HFT_{i,d}$	$QT_{i,d}$	0.89*** (18.67)	1.46*** (11.73)	1.31*** (17.34)
	$CT_{i,d}$	0.60*** (11.84)	1.35*** (11.54)	0.96*** (11.58)
	$OR_{i,d}$	2.76*** (32.03)	1.67*** (5.86)	3.59*** (27.21)
	$OV_{i,d}$	1.89*** (25.48)	1.50*** (6.53)	2.61*** (24.75)
	$ITS_{i,d}$	1.13*** (12.79)	1.54*** (5.06)	2.53*** (19.68)
	Controls	Yes	Yes	Yes
Time and Stock FEs		Yes	Yes	Yes
N		2,969,829	640,306	2,967,095

**Table A.2. First stage instrumental variable (IV) regression results – MIDAS sample**

This table presents the results for the estimation of the impact of the selected instruments on HFT measures:

$$HFT_{i,d} = \alpha_i + \beta_d + \vartheta_1 IV_{i,d} + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d}$$

where  $HFT_{i,d}$  corresponds one of the five HFT proxies ( $QT_{i,d}$ ,  $CT_{i,d}$ ,  $OR_{i,d}$ ,  $OV_{i,d}$ ,  $ITS_{i,d}$ ).  $\alpha_i$  and  $\beta_d$  are stock and time (day) fixed effects. The  $C_{k,i,d}$  is a set of  $k$  control variables, including variables from both the option and underlying markets. The option market variables are  $Ovolume_{i,d}$ ,  $Oimplied_{i,d}$ ,  $|Odelta_{i,d}|$ ,  $Ogamme_{i,d}$  and  $Ovega_{i,d}$  and the stock market variables are  $SDspread_{i,d}$  and  $SVolatility_{i,d}$ . Standard errors are double clustered on stock and day. For the definitions and computation methods of all the variables, see Table 1. In Columns 1, 3, 5 and 7, the level of HFT is instrumented with the changes in tick size; for this specification,  $IV_{i,d}$  is a dummy variable that takes the value of one after the change (from October 3, 2016 to September 28, 2018), and zero before (from October 1, 2014 to October 2, 2016) for the treatment stocks and  $IV_{i,d}$  takes the value of zero in the entire period for the control stocks. In Column 2, 4, 6 and 8, the level of HFT is instrumented with the average level of HFT on that day in all other stocks in the corresponding size quintile; for this specification,  $IV_{i,d}$  is the average level of five HFT proxies ( $QT_{i,d}$ ,  $CT_{i,d}$ ,  $OR_{i,d}$ ,  $OV_{i,d}$ ,  $ITS_{i,d}$ ) in all other stocks in the corresponding size quintile. For Column 1, 3, 5 and 7, the sample contains 1,235 stocks (617 control stocks and 618 treated stocks) that implemented in the SEC's Tick Size Pilot Programme from October 1, 2014, to September 28, 2018. For Columns 2, 4, 6 and 8, the sample contains 2,746 stocks traded between January 1, 2012, and December 31, 2019, on the US exchanges. We follow Bollen and Whaley (2004) and define OTM options as those with  $|Odelta_{i,t}| \leq 0.375$ , ATM options as those with  $0.375 < |Odelta_{i,t}| \leq 0.625$ , and ITM options as those with  $|Odelta_{i,t}| > 0.625$ . Standard errors are double clustered on stock and day and t-statistics are reported in parentheses. \*, \*\* and \*\*\* denote significance at 10%, 5% and 1%.

Dependent Variable	Full sample		ATM		ITM		OTM	
	IV (Tick Size Pilot) (1)	IV (Average HFT) (2)	IV (Tick Size Pilot) (3)	IV (Average HFT) (4)	IV (Tick Size Pilot) (5)	IV (Average HFT) (6)	IV (Tick Size Pilot) (7)	IV (Average HFT) (8)
$QT_{i,t}$	-0.30*** (-98.58)	0.01*** (2.89)	-0.31*** (-83.20)	0.01*** (4.95)	-0.31*** (-74.90)	0.001* (1.83)	-0.30*** (-89.06)	0.05*** (3.78)
$CT_{i,t}$	-0.35*** (-127.24)	0.04*** (45.76)	-0.35*** (-107.74)	0.05*** (45.34)	-0.35*** (-98.34)	0.04*** (32.84)	-0.34*** (-113.69)	0.04*** (43.11)
$OR_{i,t}$	-0.09*** (-66.38)	0.38*** (307.8)	-0.08*** (-51.43)	0.36*** (269.9)	-0.08*** (-43.73)	0.36*** (250.9)	-0.08*** (-58.37)	0.36*** (284.2)
$OV_{i,t}$	-0.14*** (-85.74)	0.32*** (266.18)	-0.14*** (-68.49)	0.31*** (233.25)	-0.14*** (-61.23)	0.29*** (212.26)	-0.14*** (-76.18)	0.31*** (246.03)
$ITS_{i,t}$	-0.07*** (-63.56)	0.21*** (193.69)	-0.07*** (-49.34)	0.22*** (181.89)	-0.07*** (-46.70)	0.21*** (166.31)	-0.07*** (56.75)	0.21*** (186.26)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Stock and Time FEs	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
N	640,306	2,967,095	402,656	2,264,040	343,460	2,077,881	480,324	2,540,617

**Table A.3. The impact of HFT on the realized volatility – MIDAS Sample**

This table presents the results for the estimation of the impact of HFT on the realized volatility:

$$SRVolatility_{i,d} = \alpha_i + \beta_d + \gamma_1 \widehat{HFT}_{i,d} + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d}$$

$$HFT_{i,d} = \alpha_i + \beta_d + \vartheta_1 IV_{i,d} + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d}$$

where  $SRVolatility_{i,d}$  is computed as the variance of secondly midpoint stock returns,  $HFT_{i,d}$  corresponds one of the five HFT proxies ( $QT_{i,d}$ ,  $CT_{i,d}$ ,  $OR_{i,d}$ ,  $OV_{i,d}$ ,  $ITS_{i,d}$ ).  $\alpha_i$  and  $\beta_d$  are stock and time (day) fixed effects. The  $C_{k,i,d}$  is a set of  $k$  control variables, including variables from both the option and underlying markets. The option market variables are  $Ovolume_{i,d}$ ,  $Oimplied_{i,d}$ ,  $|Odelta_{i,d}|$ ,  $Ogamme_{i,d}$  and  $Ovega_{i,d}$  and the stock market variables are  $SPspread_{i,d}$  (when we employ  $OPspread_{i,d}$  as the dependent variable),  $SDspread_{i,d}$  (when we use  $ODspread_{i,d}$  as the dependent variable), and  $SVolatility_{i,d}$ . For the definitions and computation methods of all the variables, see Table 1. Three specifications of the model are estimated. In Column 1, we estimate the model by using OLS with stock and fixed effects. In Columns 2 and 3, we use 2SLS IV approach. In Column 2, the level of HFT is instrumented with the changes in tick size; for this specification,  $IV_{i,d}$  is a dummy variable that takes the value of one after the change (from October 3, 2016 to September 28, 2018), and zero before (from October 1, 2014 to October 2, 2016) for the treatment stocks and  $IV_{i,d}$  takes the value of zero in the entire period for the control stocks. In Column 3, the level of HFT is instrumented with the average level of HFT on that day in all other stocks in the corresponding size quintile; for this specification,  $IV_{i,d}$  is the average level of five HFT proxies ( $QT_{i,d}$ ,  $CT_{i,d}$ ,  $OR_{i,d}$ ,  $OV_{i,d}$ ,  $ITS_{i,d}$ ) in all other stocks in the corresponding size quintile. For Columns 1 and 3, the sample contains 2,746 stocks traded between January 1, 2012, and December 31, 2019, on the US exchanges. For Column 2, the sample contains 1,235 stocks (617 control stocks and 618 treated stocks) that implemented in the SEC's Tick Size Pilot Programme from October 1, 2014, to September 28, 2018. Standard errors are double clustered on stock and day and t-statistics are reported in parentheses. \*, \*\* and \*\*\* denote significance at 10%, 5% and 1%.

<i>SRVolatility<sub>i,d</sub></i> is the dependent variable.				
	Variable	OLS (1)	IV (Tick Size Pilot) (2)	IV (Average HFT) (3)
<i>HFT<sub>i,d</sub></i>	<i>QT<sub>i,d</sub></i>	-0.001*** (-45.22)	-0.002*** (-28.99)	-0.002*** (-44.14)
	<i>CT<sub>i,d</sub></i>	-0.002*** (-56.77)	-0.002*** (-28.18)	-0.003*** (-39.99)
	<i>OR<sub>i,d</sub></i>	0.01*** (121.30)	-0.01*** (-21.88)	0.02*** (173.34)
	<i>OV<sub>i,d</sub></i>	0.01*** (96.29)	-0.002*** (-20.82)	0.01** (140.18)
	<i>ITS<sub>i,d</sub></i>	0.01*** (82.27)	-0.002*** (-21.26)	0.01*** (124.06)
	Controls	Yes	Yes	Yes
	Time and Stock FEs	Yes	Yes	Yes
	N	2,969,829	640,306	2,967,095



**Table A.4. The impact of HFT on option spread after controlling for the realized volatility – MIDAS Sample**

This table presents the results for the estimation of the impact of HFT on the options spread:

$$OSpread_{i,d} = \alpha_i + \beta_d + \gamma_1 HFT_{i,d} + \sum_{k=1}^8 \delta_k C_{k,i,d} + \varepsilon_{i,d}$$

$$HFT_{i,d} = \alpha_i + \beta_d + \vartheta_1 IV_{i,d} + \sum_{k=1}^8 \delta_k C_{k,i,d} + \varepsilon_{i,d}$$

where  $OSpread_{i,d}$  corresponds to either the proportional spread ( $OPspread_{i,d}$ ) or the dollar spread ( $ODspread_{i,d}$ ),  $HFT_{i,d}$  corresponds one of the five HFT proxies ( $QT_{i,d}$ ,  $CT_{i,d}$ ,  $OR_{i,d}$ ,  $OV_{i,d}$ ,  $ITS_{i,d}$ ).  $\alpha_i$  and  $\beta_d$  are stock and time (day) fixed effects. The  $C_{k,i,d}$  is a set of  $k$  control variables, including variables from both the option and underlying markets. The option market variables are  $Ovolume_{i,d}$ ,  $Oimplied_{i,d}$ ,  $|Odelta_{i,d}|$ ,  $Ogamme_{i,d}$  and  $Ovega_{i,d}$  and the stock market variables are  $SPspread_{i,d}$  (when we employ  $OPspread_{i,d}$  as the dependent variable),  $SDspread_{i,d}$  (when we use  $ODspread_{i,d}$  as the dependent variable),  $SVolatility_{i,d}$  and quote-based realized volatility ( $SRVolatility_{i,d}$ ).  $SRVolatility_{i,d}$  is computed as the variance of secondly midpoint stock returns. For the definitions and computation methods of all the variables, see Table 1. Three specifications of the model are estimated. In Column 1, we estimate the model by using OLS with stock and fixed effects. In Columns 2 and 3, we use 2SLS IV approach. In Column 2, the level of HFT is instrumented with the changes in tick size; for this specification,  $IV_{i,d}$  is a dummy variable that takes the value of one after the change (from October 3, 2016 to September 28, 2018), and zero before (from October 1, 2014 to October 2, 2016) for the treatment stocks and  $IV_{i,d}$  takes the value of zero in the entire period for the control stocks. In Column 3, the level of HFT is instrumented with the average level of HFT on that day in all other stocks in the corresponding size quintile; for this specification,  $IV_{i,d}$  is the average level of five HFT proxies ( $QT_{i,d}$ ,  $CT_{i,d}$ ,  $OR_{i,d}$ ,  $OV_{i,d}$ ,  $ITS_{i,d}$ ) in all other stocks in the corresponding size quintile. For Columns 1 and 3, the sample contains 2,746 stocks traded between January 1, 2012, and December 31, 2019, on the US exchanges. For Column 2, the sample contains 1,235 stocks (617 control stocks and 618 treated stocks) that implemented in the SEC's Tick Size Pilot Programme from October 1, 2014, to September 28, 2018. Standard errors are double clustered on stock and day and t-statistics are reported in parentheses. \*, \*\* and \*\*\* denote significance at 10%, 5% and 1%.

Panel A: $ODspread_{i,t}$ is the dependent variable.				
	Variable	OLS (1)	IV (Tick Size Pilot) (2)	IV (Average HFT) (3)
$HFT_{i,d}$	$QT_{i,d}$	0.02*** (29.40)	0.02*** (11.03)	0.03*** (26.12)
	$CT_{i,d}$	0.01*** (16.81)	0.02*** (11.20)	0.01*** (11.43)
	$OR_{i,d}$	0.06*** (51.40)	0.02*** (6.95)	0.10*** (50.97)
	$OV_{i,d}$	0.05*** (48.67)	0.02*** (8.61)	0.08*** (52.06)
	$ITS_{i,d}$	0.05*** (35.65)	0.02*** (6.40)	0.07*** (35.39)
	Controls	Yes	Yes	Yes
	Time and Stock FEs	Yes	Yes	Yes
	N	2,969,829	640,306	2,967,095
Panel B: $OPspread_{i,t}$ is the dependent variable.				
	Variable	OLS (1)	IV (Tick Size Pilot) (2)	IV (Average HFT) (3)
$HFT_{i,d}$	$QT_{i,d}$	0.93*** (20.24)	1.48*** (11.34)	1.37*** (18.66)
	$CT_{i,d}$	0.64*** (12.87)	1.37*** (15.30)	1.04*** (12.51)
	$OR_{i,d}$	2.57*** (29.67)	1.72*** (5.96)	3.18*** (23.40)
	$OV_{i,d}$	1.76*** (25.29)	1.54*** (7.31)	2.33*** (21.74)
	$ITS_{i,d}$	0.93*** (10.00)	1.59*** (5.33)	2.18*** (16.39)
	Controls	Yes	Yes	Yes
	Time and Stock FEs	Yes	Yes	Yes
	N	2,969,829	640,306	2,967,095

**Table A.5. Summary statistics for NASDAQ Sample**

This table reports the descriptive statistics for the variables used in our analysis. Panel A provides the descriptive statistics for all options-related variables separately for the full sample and three groups based on moneyness. Panel B shows the descriptive statistics for all variables from the underlying stock market. For the definitions and computation methods of the variables, see Table 1. We follow Bollen and Whaley (2004) and define OTM options as those with absolute option delta  $|\Delta| \leq 0.375$ , ATM options as those with  $0.375 < |\Delta| \leq 0.625$ , and ITM options as those with  $|\Delta| > 0.625$ . The sample contains 103 stocks traded between January 1, 2009, and December 31, 2009, on the NASDAQ.

Panel A. Equity market variables				
	Variable	Mean	Median	Stdev
Full Sample	$SHFT_{i,d}^{All}$	0.49	0.48	0.21
	$SHFT_{i,d}^D$	0.33	0.31	0.16
	$SHFT_{i,d}^S$	0.25	0.25	0.17
	$SPspread_{i,d}$	0.11	0.12	0.23
	$SDspread_{i,d}$ (\$)	0.03	0.02	0.04
	$SVolatility_{i,d}$	0.41	0.35	6.65
	$SOIB_{i,d}$	0.09	0.07	0.09
Panel B. Option market variables				
Full sample	$ODspread_{i,t}$	0.136	0.100	0.137
	$OPspread_{i,t}$ (%)	5.61	4.52	5.25
	$Ovolume_{i,t}$	5.892	6.276	2.762
	$Oimplied_{i,t}$	0.407	0.359	0.245
	$ Odelta_{i,t} $	0.539	0.530	0.247
	$Ogamma_{i,t}$	0.11	0.09	0.14
	$Ovega_{i,t}$	2.403	1.910	2.330
ATM	$ODspread_{i,t}$	0.144	0.117	0.178
	$OPspread_{i,t}$ (%)	0.056	0.046	0.089
	$Ovolume_{i,t}$	6.900	7.154	2.455
	$Oimplied_{i,t}$	0.398	0.344	0.244
	$ Odelta_{i,t} $	0.476	0.469	0.159
	$Ogamma_{i,t}$	0.13	0.09	0.16
	$Ovega_{i,t}$	4.041	3.505	3.736
ITM	$ODspread_{i,t}$	0.207	0.163	0.216
	$OPspread_{i,t}$ (%)	0.037	0.031	0.073
	$Ovolume_{i,t}$	4.295	4.444	2.467
	$Oimplied_{i,t}$	0.502	0.468	0.286
	$ Odelta_{i,t} $	0.732	0.703	0.182
	$Ogamma_{i,t}$	0.10	0.09	0.18
	$Ovega_{i,t}$	2.308	2.001	2.316
OTM	$ODspread_{i,t}$	0.103	0.098	0.134
	$OPspread_{i,t}$ (%)	0.075	0.069	0.091
	$Ovolume_{i,t}$	5.331	5.578	2.499
	$Oimplied_{i,t}$	0.206	0.202	0.202
	$ Odelta_{i,t} $	0.110	0.107	0.088
	$Ogamma_{i,t}$	0.08	0.07	0.12
	$Ovega_{i,t}$	2.193	1.691	2.478

**Table A.6. First stage instrumental variable (IV) regression results – NASDAQ sample.**

This table presents the results for the estimation of the impact of the selected instruments on HFT measures:

$$SHFT_{i,d}^D = \alpha_i + \beta_d + \vartheta_1 IV_{i,d} + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d}$$

$$SHFT_{i,d}^S = \alpha_i + \beta_d + \vartheta_1 IV_{i,d} + \sum_{k=1}^7 \delta_k C_{k,i,d} + \varepsilon_{i,d}$$

where  $SHFT_{i,d}^D$  and  $SHFT_{i,d}^S$  are the measures of HFTs-liquidity demanding and supplying activities, respectively.  $\alpha_i$  and  $\beta_d$  are stock and time (day) fixed effects. The  $C_{k,i,d}$  is a set of  $k$  control variables, including variables from both the option and underlying markets. The option market variables are  $Ovolume_{i,d}$ ,  $Oimplied_{i,d}$ ,  $|Odelta_{i,d}|$ ,  $Ogamme_{i,d}$  and  $Ovega_{i,d}$  and the stock market variables are  $SDspread_{i,d}$  and  $SVolatility_{i,d}$ . Standard errors are double clustered on stock and day. For the definitions and computation methods of all the variables, see Table 1. Two specifications of the model are estimated. In Columns 1 and 3,  $IV_{i,d}$  is a dummy variable that takes the value 1 during the flash-orders period (from June 5, 2009, to August 31, 2009) initiated by the NASDAQ. In Columns 2 and 4, the level of HFT is instrumented with the average level of HFT on that day in all other stocks in the corresponding size quintile; for this specification,  $IV_{i,d}$  is the average level of two HFT proxies ( $SHFT_{i,d}^D$  and  $SHFT_{i,d}^S$ ) in all other stocks in the corresponding size quintile. The sample contains 103 stocks traded between January 1, 2009, and December 31, 2009, on the NASDAQ. Standard errors are double clustered on stock and day, and t-statistics are reported in parentheses. \*, \*\* and \*\*\* denote significance at 10%, 5% and 1%.

	$SHFT_{i,d}^{All}$		$SHFT_{i,d}^D$		$SHFT_{i,d}^S$	
	IV (Flash orders) (1)	IV (Average HFT) (2)	IV (Flash orders) (3)	IV (Average HFT) (4)	IV (Flash orders) (5)	IV (Average HFT) (6)
$IV_{i,d}$	3.52*** (2.91)	0.27*** (9.89)	3.12*** (2.70)	0.20*** (8.37)	2.24** (2.52)	0.18*** (7.21)
Controls	Yes	Yes	Yes	Yes	Yes	Yes
Stock FE	Yes	Yes	Yes	Yes	Yes	Yes
Time FE	No	Yes	No	Yes	No	Yes